Food Security, Climate Change and Environmental Pollution in MENA Region: Evidence from Second Generation Panel Analysis

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ABSTRACT

Food security is a critical issue in the Middle East and North Africa (MENA) region due to its population growth, as well as geographical and climatic conditions. From one point of view, most of the countries in the region benefit from an abundance of natural resources centered on fossil fuels. From another point of view, environmental issues, particularly emissions caused by production activities, and the pressures caused by climate variability, highlight the importance of food security. Hence, the effects of climate change, energy consumption, environmental pollution and other control variables on food security in the MENA region were explored from 1990 to 2019. According to the crosssection dependency, the second-generation panel CS-ARDL (Cross-Sectional Autoregressive Distributed Lag) estimator was employed. The empirical results indicate that energy consumption, crop production land, CO₂ emissions, and precipitation have a significant positive effect on crop production index, as index of food security. Additionally, urbanization and mean temperature have detrimental effects. The findings from Dumitrescu and Hurlin causality tests indicated that crop land and precipitation have a unidirectional causal effect on food security, whereas energy consumption, CO₂ emissions, urbanization, and mean temperature have a bidirectional causal relationship with food security. These findings imply that while maintaining the level of agricultural production and increasing it, the climate effects and environmental aspects of production should not be overlooked.

Keywords: CO2 emissions, CS-ARDL, Energy consumption.

INTRODUCTION

The Second Sustainable Development Goal (SDG2), has set the target of enhancing nutrition, attaining food security, eradicating hunger, and promoting sustainable agriculture by the year 2030. Conflict, climate variability, and economic downturns have hindered progress toward SDG2 over the last few years, and these factors are expected to worsen following COVID-19, which is now being exacerbated by the Ukraine-Russia crisis. Between 720 and 811 million people worldwide go to bed hungry every night, highlighting the serious consequences of the current global crises (UNICEF, 2020). Moreover, the number of people experiencing extreme food insecurity has doubled since COVID-19, increasing from 135 million to 276 million (UN Secretary General, 2022). Following the World Health Organization, the likelihood of becoming undernourished increased to 9.9% in 2020 from 8.4% in 2019 (WHO,

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In 1996, the World Food Summit stated that food security is achieved when every individual has access to sufficient and safe food supply that sustains an active and healthy life (World Food Summit, 1996). In this regard, the Food and Agriculture Organization identifies (FAO) four fundamental dimensions of food security: physical food availability, food access, food utilization, and food stability (Webb et al., 2006; CFS, 2009). Physical food availability is achieved when a sufficient amount of food is permanently available for all members of the society. In this dimension of food security, water, land and energy use determine the food production growth (Godfray et al., 2010). The agricultural sector plays a key role in this dimension of food security. Since the dawn of humanity, agriculture has provided food for humans and contributed to the improvement of human living standards.

While global institutions such as FAO, WFP (World Food Programme), and the IFAD (International Fund for Agricultural Development) play a significant role in achieving the second SDG: domestic strategies, such as increasing agricultural productivity and promoting sustainable food production, which are the most effective means of achieving food security and global zero hunger. The increasing global population, projected to reach 11.2 billion by 2100, is driving a rising demand for food and agricultural products. As the population continues to grow and food production rises, it is imperative to prioritize and increase production agricultural to fulfill the increasing demand for food of human societies. Several recent studies, such as Lu et al. (2021), predicted that given the current consumption patterns, food, water, and energy consumption would rise by 50%, 80%, and 60%, respectively, for a population of 10 billion, by 2050. A variety of factors, including land degradation, water scarcity, and global warming, are threatening food production. To feed 11.2 billion people by 2100, global food production needs to rise more than 50%. Increased food production will also pose numerous environmental challenges (Searchinger *et al.*, 2019).

Extensive research has explored the interplay of food security with various factors. climate including change (Schmidhuber and Tubiello, 2007; Campbell et al., 2017; Mokhtar et al., 2022; Pickson and Boateng, 2022; Kargar Dehbidi et al., 2022), CO₂ emissions (Chandio et al., 2020; Degife et al., 2021; Koondhar et al., 2021a; Affoh et al., 2022), fossil fuel consumption (Günther, 2001; Arizpe et al., 2011; Raeeni et al., 2019; Mahdavian et al., 2022; Boly and Sanou, 2022), renewable energy consumption (Mallick, 2022; Kaimal et al., 2022), population (Rehman et al., 2022), economic growth (Kargar Dehbidi et al., 2022), water resources (Abdullah et al., 2022), soil fertility (Gebrehiwot, 2022), agricultural land (Hossain et al., 2020), environmental deterioration (Qi et al., 2018), and urbanization (Wang, 2019) across diverse countries and regions. This research has employed variety of econometric а techniques and methods.

Schmidhuber and Tubiello (2007) studied the impact of climate change on four dimensions of food security, finding a detrimental effect on all aspects. They noted that climate changes overall affect food security, and is regionally and temporally variable, contingent upon a country's socioeconomic status when addressing climate change. Raeeni et al. (2019) employed time series econometric methods, including causality and co-integration tests, confirming significant relationship among energy consumption and agricultural products in Iran.

Also, Kargar Dehbidi *et al.* (2022) examined the effect of climate change (precipitation and temperature) on food security (food price volatility) in Iran's provinces, utilizing the Panel-Var econometrics approach. Empirical findings revealed a significant effect of climate change on food security, with temperature exerting a greater influence than precipitation. Onour (2019) employed the ARDL bounds test of co-integration to assess CO₂ emissions' impact on Sudan's crop yields, revealing a significant positive impact on cereal yield. A 1% increase in CO₂ emissions resulted in 3% and 0.7% increase in cereal yield in short and long run, respectively, a finding echoed by Degife et al. (2021) for maize yields in Ethiopia. Affoh et al. (2022) investigated CO₂ emissions' impact on food security subindices (food availability, accessibility, and utilization) using PMG, FMOLS, and DOLS models across 25 sub-Saharan African nations. They found that CO₂ emissions had no significant impact on food utilization, but had a positive impact on food accessibility and availability. Regarding the effect of energy consumption on agricultural products, Numerous studies have looked at how CO₂ emissions in MENA nations are impacted by factors like energy use, crop production, and urbanization (Farhani and Rejeb, 2012; Arouri et al., 2012; Omri, 2013; Jebli and Youssef, 2017; Magazzino and Cerulli, 2019; Alharthi et al., 2021; Omri and Saidi, 2022).

Nonetheless, according to the authors' analysis, there has not been a comprehensive study conducted that analyzes the impact of CO_2 emissions on crop production index within this region. Identifying this existing research gap highlights the necessity and significance of this research as follows. First, even with the evident importance of CO_2 emissions and other control variables in influencing crop yields, a comprehensive investigation spanning the MENA region has not been undertaken. By addressing this gap, the study contributes to a deeper understanding of the dynamics of food

security in MENA countries. The present study's findings will elucidate the primary determinants of food insecurity, providing valuable insights for the achievement of SDG, particularly within the MENA region. Secondly, this study pioneers the examination of the food security-energyclimate change nexus in the MENA context, thus enhancing comprehension of the intricate challenges faced by MENA nations. the study delves into Thirdly, the relationships among CO₂ emissions, fossil fuel consumption, cropland, urbanization, temperature, precipitation, and crop production as a food security indicator. This exploration is conducted using the secondgeneration panel CS-ARDL estimator across a panel of 18 MENA countries. Lastly, the integration of recent methodological advancements, including second-generation panel tests, further bolsters the study's findings, enhancing their robustness and accuracy.

MATERIALS AND METHODS

Data

According to the empirical study's goals and data availability, the data was collected from 1990 to 2019 for 18 MENA countries. Table 1 illustrates the details of variables of econometrics model.

Crop Production index (CP), CO₂ emission (CO₂), Crop Land (CPL), and Mean Temperature (MT) were collected from the FAO. Urban Population (URB), and

Table 1. Details of the model's variables.

| Variables | Definition | Unit of measurement |
|--|---|--------------------------------|
| Crop Production index (CP) | All agricultural production (except fodder) relative to the base period $(2014-2016 = 100)$ | Unit less (Index) |
| Cropland (CPL) | Land used for the cultivation of crops | 1000 ha |
| Urban Population (URB) | The share of urban to total population | Percent |
| Energy Consumption (EC) | Total energy consumption | Million tons of oil equivalent |
| CO ₂ Emissions (CO ₂) | Total CO ₂ emissions by agri-food system component | Kilotons |
| Mean Temperature (MT) | Annual Mean Temperature | Centigrade |
| Precipitation (PRC) | Annual Mean Precipitation | Millimeter |

Precipitation (PRC) were gathered from the World Bank. Also, the Energy Consumption (EC) data obtained from Energy Information Administration (EIA).

Model and Econometrics Method

According to the literature, the variables of model are selected. Hence, the empirical econometrics model can be expressed by Equation (1):

 $CP_{it} = \alpha_0 + \alpha_1 CPL_{it} + \alpha_2 URB_{it} + \alpha_3 EC_{it} + \alpha_4 CO2_{it} + \alpha_5 MT_{it} + \alpha_6 PRC_{it} + \varepsilon_{it}$ (1)

Equation (2) indicates the ARDL approach, while the expanded form of Equation (1) is shown in Equation (3), taking into account the cross-sectional averages of the variables in the studied model (Chudik and Pesaran, 2015; Shao *et al.*, 2021; Chien *et al.*, 2022).

$$W_{i,t} = \sum_{i=1}^{P_{w}} \vartheta_{i,t} W_{i,t-1} + \sum_{i=0}^{P_{x}} \rho_{i,t} X_{i,t-1} + \varepsilon_{i,t}^{\prime}$$
(2)

$$W_{i,t} = \sum_{i=1}^{N} \vartheta_{i,t} W_{i,t-1} + \sum_{i=0}^{N} \rho_{i,t} X_{i,t-1} + \sum_{i=0}^{P_z} \beta_{i,t} \bar{Z}_{t-1} + \varepsilon_{i,t}$$
(3)

Where, i denote the cross-section (18 MENA region countries) and t denotes time period (1990 to 2019). W_{it} and X_{i, t-1} indicate the dependent and independent variables, respectively. Additionally, \bar{Z}_{t-1} represents the average of sections to address cross-sectional dependence. P_w, P_x, and P_z, imply the lags. For the long-term estimation using CS-ARDL, the average mean group estimate is presented in Equation (4). The short-term model is revealed in Equation (5) as follows: (Adebayo *et al.*, 2023; Li *et al.*, 2023).

$$\begin{aligned} \hat{\pi}_{CS-ARDL,i} &= \frac{\sum_{l=0}^{P_{x}} \hat{\rho}_{li}}{1 - \sum_{l=0} \hat{\theta}_{l,t}} \end{aligned} \tag{4} \\ \Delta W_{it} &= \varphi_{i} \Big[W_{i,t-1} - \pi_{i} X_{i,t-1} \Big] \\ &- \sum_{\substack{P_{w} \\ P_{x}}} \theta_{i,t} \Delta_{i} W_{i,t-1} \\ &+ \sum_{\substack{P_{x} \\ P_{x}}} \rho_{i,t} \Delta_{i} X_{i,t-1} \\ &+ \sum_{\substack{i=0 \\ P_{x}}} \beta_{i} \bar{Z}_{t} + \varepsilon_{i,t} \end{aligned}$$

(5)

Furthermore, all variables in the model, except for urbanization (percent), were converted to natural logarithms to reduce scale differences and improve estimation efficiency. Finally, the CS-ARDL equation for the variables in the present study is as follows:

$$\begin{split} \Delta lnCP_{i,t} &= \theta_i + \sum_{i=1}^{p} \theta_{i,t} \Delta lnCP_{i,t-1} + \\ \sum_{i=1}^{p} \theta_{i,t} \Delta lnCPL_{i,t} \sum_{i=1}^{p} \theta_{i,t} \Delta lnURB_{i,t} + \\ \sum_{i=1}^{p} \theta_{i,t} \Delta lnEC_{i,t} + \sum_{i=1}^{p} \theta_{i,t} \Delta lnCO2_{i,t} + \\ \sum_{i=1}^{p} \theta_{i,t} \Delta lnMT_{i,t} + \sum_{i=1}^{p} \theta_{i,t} \Delta lnPRC_{i,t} + \\ \sum_{i=0}^{p} \beta_{i,t} \bar{\mathcal{L}}_{i,t-1} + \varepsilon_{i,t} \end{split}$$
(6)

Initially, cross-sectional dependency should be checked in the empirical panel data. Therefore, the Pesaran (2004) crosssection test (Pesaran CD test) is applied to examine the presence of cross-sectional dependency for all variables in the model. In the Pesaran CD test, the null hypothesis is the absence of cross-section dependence (Pesaran *et al.*, 2008). Equation (7) presents the Pesaran CD test statistic (Pesaran, 2004).

$$CD = \sqrt{\frac{2T}{N(N-1)}} \sum_{i=1}^{N-1} \sum_{j=i+1}^{N} \hat{\rho}_{ik}$$
(7)

Where, T is the time period (20 years) and N denotes the cross-section (18 MENA countries). Additionally, $\hat{\rho}_{ij}$ represents the correlation coefficient. According to the results of Pesaran CD test, the researchers could select the first or second generation unit root tests. The first generation unit root tests contain Levin, Lin and Chu (LLC) and Im, Pesaran and Shin (IPS) stationary test. The second-generation unit root tests contain Cross-Sectionally Augmented IPS (CIPS) stationary test.

However, it is necessary to check the homogeneity of slope in all cross-sections, before estimating the econometric model. According to this, the Pesaran and Yamagata (2008) homogeneity test was used in the present study. The null and alternative hypothesis of the slope homogeneity test is homogenous and heterogeneous slopes of cross-section, respectively (Pesaran and Yamagata, 2008). The homogeneity of slope is checked by Equations (8) and (9).

$$\tilde{\Delta} = \sqrt{N} \left(\frac{N^{-1} S\% - k}{\sqrt{2k}} \right) \tag{8}$$

$$\tilde{\Delta}_{\text{Adjusted}} = \sqrt{N} \left(\frac{N^{-1} S\% - k}{\sqrt{\frac{2k(T - k - 1)}{T + 1}}} \right)$$
(9)

In this study, the Westerlund panel cointegration test as the second-generation cointegration test is used to select the appropriate econometrics estimation approach. Following Westerlund (2007), the panel co-integration is checked by Equations (10) to (13).

$$G_a = \frac{1}{n} \sum_{i=1}^{n} \frac{\dot{\alpha}_i}{SE(\dot{\alpha}_i)} \tag{10}$$

$$G_t = \frac{1}{n} \sum_{i=1}^{n} \frac{T \dot{\alpha}_i}{\dot{\alpha}_i(1)} \tag{11}$$

$$P_a = T \, \acute{\propto} \tag{12}$$

$$P_t = \frac{\alpha}{SE(\alpha)} \tag{13}$$

In this paper, the second-generation panel CS-ARDL estimator is utilized because of its advantages over other methods. Panel CS-ARDL provides robust, effective, and powerful estimation capabilities, even in the presence of non-stationarity, slope misspecification heterogeneity, bias. endogeneity bias, serial correlation of error terms, limited sample size, and crosssectional dependency (Samargandi, 2019; Azam and Haseeb, 2021; Okunade et al., 2022; and Salman *et* al., 2022). Additionally, CS-ARDL can estimate both long and short-run relationships, simultaneously. Moreover, the lag of dependent and independent variables can be included in the econometric model (Chudik and Pesaran, 2015)

RESULTS

The descriptive statistics of all variables of the model is showed in Table 2.

According to the results of Table 2, the mean of LnCP is 4.42, while the mean of $LnCO_2$ is 8.96. Furthermore, the mean of LnEC is 3.08, whereas the mean of LnCPL is 6.55. Also, the mean of URB, LnMT, and LnPRC are 72.4, 3.1, and 4.8, respectively, in the MENA region. The highest values of standard deviation belong to the LnURB and the lowest values to the LnMT variable.

As mentioned before, the cross-section dependence of variables must be checked before the stationary test (Westerlund, 2007; Salim *et al.*, 2017; Shao *et al.*, 2021; Tarazkar, *et al.*, 2021; Chien *et al.*, 2022). The results of Pesaran CD test are reported in Table 3.

The results of the Pesaran CD test strongly rejected the null hypothesis of no crosssection dependence for all variables in the model, except for LnCPL. Since all variables (except LnCPL) exhibit crosssectional dependence, it is recommended to use the second-generation panel stationary test. Therefore, the CIPS panel stationary test was employed to check the stationary properties of all variables, except LCPL. In conformity with the results of the Pesaran CD test, the LLC and IPS tests were used for LnCPL. The results of the CIPS, IPS, and LLC panel stationary tests are presented in Table 4

Table 2. Descriptive statistics of variables for MENA countries.

| Variables | LnCP | LnCPL | URB | LnEC | LnCO ₂ | LnMT | LnPRC |
|--------------------|-------|-------|-------|------|-------------------|-------|-------|
| Mean | 4.42 | 6.55 | 72.4 | 3.08 | 8.96 | 3.1 | 4.8 |
| Median | 4.49 | 7.54 | 76.02 | 2.96 | 8.84 | 3.13 | 4.75 |
| Maximum | 5.58 | 9.83 | 100 | 5.71 | 11.8 | 3.37 | 6.81 |
| Country | UAE | IRI | KWT | IRI | IRI | BHR | LBN |
| Minimum | 1.72 | 1.38 | 20.93 | 0.58 | 6.6 | 2.64 | 2.63 |
| Country | KWT | BHR | YEM | MAR | YEM | LBN | QAT |
| Standard Deviation | 0.41 | 2.47 | 18.7 | 1.08 | 1.15 | 0.16 | 0.8 |
| Skewness | -1.55 | -0.64 | -0.64 | 0.45 | 0.48 | -0.32 | 0.07 |
| Kurtosis | 9.94 | 2.17 | 2.8 | 2.56 | 2.55 | 2.12 | 2.47 |
| Observations | 540 | 540 | 540 | 540 | 540 | 540 | 540 |
| Cross section | 18 | 18 | 18 | 18 | 18 | 18 | 18 |

| 17 | | D |
|----------|-------------------|-------------|
| Table 3. | Results of Pesara | in CD test. |

| Variables | Pesaran CD test |
|-------------------|-----------------|
| LnCP | 23.95*** |
| LnCPL | 0.801 |
| URB | 49.28*** |
| LnEC | 47.34*** |
| LnCO ₂ | 36.48*** |
| LnMT | 53.52*** |
| LnPRC | 14.14*** |

*** denote significance levels at 1%

Based on the findings of Table 4, the CIPS test statistics for LnCP, LnMT, and LnPRC are statistically significant at the 1% and 5%, respectively. This suggests that LnCP, LnMT, and LnPRC follow an I(0) process. In contrast, the null hypothesis of stationary is rejected for LnCO₂, LnEC, and URB at the level (I(0)). Additionally, the CIPS test statistics for the first difference of LnCO₂, LnEC, and URB are statistically significant at the 1% level of significance. Hence, $LnCO_2$, LnEC, and URB follow an I(1) process. According to the last row of Table 4, the LLC and IPS stationary tests' statistics indicate that LnCPL is stationary at the level and follows an I(0) process. Therefore, all variables in the model follow either an I(1)or I(0) process, and none of them follow an I(2) process. In the next step, we investigated the slope homogeneity analysis. The results of the homogeneity test are presented in Table 5.

According to both $\tilde{\Delta}$ and $\tilde{\Delta}_{Adjusted}$ tests, the null hypothesis of homogenous slope

parameters is rejected at a 1% significance level, indicating the presence of slope heterogeneity across MENA countries. The results of the slope homogeneity test recommend the use of a heterogeneous econometric panel regression method. In the next step, panel cointegration tests were conducted. Table 6 showed the results of the Westerlund panel co-integration test.

The results from Table 6 confirm the presence of a long-run co-integration relationship. Therefore, the CS-ARDL approach was employed to examine the impact of independent variables on food security. The results of short and long run second-generation panel analysis are presented in Table 7.

The empirical findings from CS-ARDL estimation presented that CO₂ was positively linked with the crop production as index of food security, in both short and long run. The positive effect of CO_2 emissions on crop production is reported in some previous studies like Weyant et al. (2018), Onour (2019), Chandio et al. (2020), Koondhar et al. (2021a), and Affoh et al. (2022). The main reason for the positive impact of CO_2 on crop production is the positive effect of CO_2 emissions in the atmosphere on photosynthesis process and crop yield. Indeed, a 1% increase in the CO_2 emission can increase crop production by 0.34% in the long run.

Also, crop production and energy consumption have a significant positive

| Variables | CIPS test statistic (Level) | CIPS test statistic (First Differences) | Result |
|-------------------|-----------------------------|---|--------|
| LnCP | -2.35*** | - | I(0) |
| URB | -1.65 | -2.16*** | I(1) |
| LnEC | -1.19 | -2.21*** | I(1) |
| LnCO ₂ | -1.56 | -1.94*** | I(1) |
| LnMT | -2.41*** | - | I(0) |
| LnPRC | -2.9*** | - | I(0) |
| Variable | LLC test statistic (Level) | IPS test statistic (Level) | Result |
| LnCPL | -3.69*** | -2.82*** | I(0) |

Table 4. Results of first and second generation unit root tests.

***, **, * denote significance levels at 1%, 5% and 10%, respectively. Schwarz-Bayesian Information Criterion (SIC) has been used for optimal lag length selection.

| Table 5. | Results | of l | Pesaran | and | Yamagata |
|------------|---------|------|-----------|-----|----------|
| (2008) slo | pe homo | gene | ity test. | | |

| Test-Statistic | Value | Prob. |
|----------------------------|----------|-------|
| $\tilde{\Delta}$ | 13.99*** | 0.00 |
| $	ilde{\Delta}_{Adjusted}$ | 16.34*** | 0.00 |

*** denotes significance levels at 1%.

| Table | 6. | Panel | cointegration | test |
|-----------|------|-------|---------------|------|
| (Westerlu | nd). | | | |

| Statistic | Value |
|-----------|-----------|
| Gt | -3.483*** |
| Ga | -8.179 |
| Pt | -12.631** |
| Ра | -11.99 |

Notes: ***, **, and * Significant levels at 1, 5 and 10% respectively.

relationship in the short and long run. The positive correlation among food security and energy consumption is consistent with Raeeni *et al.* (2019), and Mahdavian *et al.* (2022). According to the long run coefficient, a 1% rise in energy consumption can boost the crop production by 0.77%. The

Table 7. Results of panel CS-ARDL estimation.

direct relationship between energy and food security implies that the higher consumption of energy leads to more crops production. Most agricultural tools and equipment are powered by fossil fuels (Ur Rahman et al., 2019). Energy in the agricultural sector is mainly used for supplying energy to water motor pumps, green house equipment, and agricultural machinery. Also, energy is used in the production process of intermediate inputs, such as fertilizers, pesticides, etc. (Martinho, 2020). Therefore, in order to increase the amount of agricultural crops, it is needed to use more agricultural equipment, which leads to increase in energy consumption.

The linkage between cropland and crop production is significant and positive: with a 1% growth in cropland, crop production rises by 0.72%. This result is consistent with Nasrullah *et al.* (2021), Koondhar *et al.* (2021b), and Kargar Dehbidi *et al.* (2022). The negative link between urbanization and crop production is not statistically significant. The effect of climate change on crop production is survived by mean

| Dependent Variable: LnCP | Coefficient | Standard error | t-Statistics |
|--------------------------|--------------|----------------|--------------|
| Long-run Results | | | |
| LnCPL | 0.72^* | 0.41 | 1.74 |
| LnURB | -0.06 | 0.129 | -0.52 |
| LnEC | 0.77^{***} | 0.28 | 2.69 |
| LnCO ₂ | 0.34** | 0.14 | 2.4 |
| LnMT | -4.58^{*} | 2.71 | -1.69 |
| LnPRC | 0.21^{*} | 0.127 | 1.69 |
| CSD-Statistics | | | -0.47 |
| Short-run Results | | | |
| $\Delta LnCP(-1)$ | 0.07 | 0.08 | 0.82 |
| ΔLnCPL | -0.14 | 0.31 | -0.47 |
| ΔLnURB | -0.04 | 0.1 | -0.43 |
| ΔLnEC | 0.55^{***} | 0.15 | 3.6 |
| $\Delta LnCO_2$ | 0.23** | 0.09 | 2.48 |
| ΔLnMT | -1.97** | 0.88 | -2.24 |
| ΔLnPRC | 0.16^{**} | 0.07 | 2.33 |
| $\Delta \text{LnEC}(-1)$ | 0.02 | 0.2 | 0.1 |
| ΔLnCPL (-1) | 0.64 | 0.43 | 1.48 |
| ΔLnPRC (-1) | 0.09^* | 0.04 | 1.92 |
| ECM (-1) | -0.92*** | 0.08 | -10.31 |

***, **, and * Significant levels at 1, 5 and 10% respectively.

temperature and precipitation. The positive effect of precipitation on crop production is statistically significant in the short and long run. This result is in line with research by Kumar et al. (2021), Ogundari and Onyaeghala (2021), and Kargar Dehbidi et al. (2022). Hence, a 1% increase in precipitation causes 0.21% increase in crop The estimated coefficient production. implies that with the rise in rainfall, the amount of available water resource boost and leads to higher production. In contrast, the temperature has a significant negative influence on crop production. Indeed, a 1% rise in temperature leads to 4.58% decline in production. It is in line with Meshram et al. (2020), and Zhang et al. (2022). Higher temperatures can increase crop growth period and evapotranspiration and reduce water availability. In general, the negative impacts of climate change primarily stem from elevated temperatures, heightened rates of evaporation and transpiration, as well as alterations in precipitation patterns, all of which have detrimental effects on crop growth. The results of the Dumitrescu and Hurlin panel causality test are reported in Table 8.

The empirical results from the employed causality tests revealed bidirectional causality between crop production (as an index of food security) and CO₂. It also

established bidirectional causality between energy use and crop production. Table 8 unidirectional causality reveals from cropland to crop production and a two-way causality link between urbanization and crop production. The findings indicate a relationship unidirectional causal from precipitation to crop production, while a bidirectional causal relationship exists between mean temperature and crop production.

DISCUSSION

Food security is one of the most essential multi-dimensional phenomena, consisting of food availability, food access, food utilization, and food stability. As a result, paying special attention to agriculture is one of the most important ways to improve food security. This sector has the most important role in the production and food security. Hence, in the present study, the factors affecting agricultural production as an index of food security are examined in the MENA countries. The dependent variable of the econometric model is the crop production index. Also, the independent variables CO_2 emission, cropland, contain precipitation, mean temperature, urban population, and energy consumption. The

| Hypothesis | W-stat | Z-stat | Results |
|-----------------------|--------------|--------|-----------------------|
| $CP \rightarrow CO_2$ | 2.06*** | 3.2 | $CP \rightarrow CO_2$ |
| $CO_2 \rightarrow CP$ | 6.49*** | 16.49 | $CO_2 \rightarrow CP$ |
| $CP \rightarrow CPL$ | 1.33 | 1.01 | $CP \rightarrow CPL$ |
| $CPL \rightarrow CP$ | 4.19*** | 9.56 | $CPL \rightarrow CP$ |
| $CP \rightarrow EC$ | 8.07^{***} | 21.23 | $CP \rightarrow EC$ |
| $EC \rightarrow CP$ | 2.33*** | 4.00 | $EC \rightarrow CP$ |
| $CP \rightarrow PRC$ | 1.13 | 0.41 | CP→PRC |
| $PRC \rightarrow CP$ | 2.02*** | 3.06 | $PRC \rightarrow CP$ |
| $CP \rightarrow MT$ | 2.11*** | 3.33 | $CP \rightarrow MT$ |
| $MT \rightarrow CP$ | 7.67*** | 20.02 | $MT \rightarrow CP$ |
| $CP \rightarrow URB$ | 6.36*** | 16.09 | $CP \rightarrow URB$ |
| $URB \rightarrow CP$ | 6.59*** | 16.77 | $URB \rightarrow CP$ |

Table 8. Results of Dumitrescu and Hurlin panel causality test.

*** Denotes significance levels at 1%.

CS-ARDL model was used to analyze panel data for the MENA countries from 1990 to 2019.

CONCLUSIONS

The outcomes of the CS-ARDL approach implied that CO₂ was positively linked with the CP in the short and long run. This finding aligns with prior studies, including those by Weyant et al. (2018), Onour (2019), Chandio et al. (2020), Koondhar et al. (2021a), and Affoh et al. (2022). The linkage between Crop Production (CP) and Energy Consumption (EC) is positive in both short and long run, which is consistent with Raeeni et al. (2019), and Mahdavian et al. (2022). This result revealed that rising energy consumption can build up crop production. Cropland directly affects production, so, expanding the CPL will lead to a rise in production costs. This result aligns with the findings of Nasrullah et al. (2021) in South Korea, Koondhar et al. (2021b) in Pakistan, and Kargar Dehbidi et al. (2022) in Iran. The association between urbanization and crop production was insignificant. Also, the effect of Temperature (MT) and Precipitation (PRC) as climatic variables on production was negative and positive, respectively, which is in line with the findings of Kumar et al. (2021), Ogundari and Onyaeghala (2021), and Kargar Dehbidi et al. (2022). The causality outcomes indicated a bidirectional causality between Crop Production (CP) and CO₂, between Energy Consumption (EC) and CP, and between Urbanization (URB) and CP. Finally, the results implied that there is a one-way causality from Precipitation (PRC) to Crop Production (CP), but the causality linkage between Mean Temperature (MT) and CP is bidirectional.

According to the empirical findings, policies must be implemented in order to create a production structure that is resistant to climate change, with a focus on minimizing pollution caused by input consumption in agricultural sectors and maintaining the foundations of sustainable development. For example, MENA countries should adopt climate-resilient agricultural practices to strengthen their farms against climate changes. They can grow drought-resistant crop varieties, practice agroforestry, and use innovative irrigation methods like drip irrigation.

Given that a substantial portion of pollution stemming from agricultural production is associated with energy consumption, the adoption of renewable energy sources, such as solar or wind power, for agricultural activities can markedly decrease carbon emissions attributed to energy use. Governments can facilitate this transition by offering financial incentives or subsidies for adopting renewable energy technologies.

Instead of chemical fertilizers and pesticides, using organic fertilizers and making producers aware of the benefits of using it is considered a suitable solution. Considering incentive policies such as guaranteed purchase of organic products, granting facilities to improve production infrastructure and imposing export subsidies on products that are produced with minimal emission of pollution and consumption of inputs can have positive effects on the production situation and food security.

Increasing the mechanization of the production sector in the studied countries can also help to minimize post-harvest losses and enhance overall productivity. Processing and packaging agricultural products can not only reduce waste, but also provide farmers with economic opportunities.

In order to lessen the negative effects of climate change and enhance food security, cultivation patterns must be tailored to the geographical conditions of each region such as drought-resistant crops in arid regions or flood-resistant varieties in areas prone to heavy rainfall.

Also, creating a communication and commercial network based on comparative advantage, available water resources and

climatic conditions can lead to increasing production stability, food security, and reducing the effects of climate change. Collaborations between governments, private sector stakeholders, and research institutions can also drive innovation and promote sustainable agricultural practices.

The current study provides valuable insights into the factors affecting food security and agricultural production in the MENA region. However, due to limited data availability, it leaves a gap in testing the impact of climate change adaptation strategies, such as drip irrigation, conservation tillage, and various livelihood activities, on food security. Investigating the effectiveness of these strategies is crucial, as they offer practical approaches to mitigate the adverse effects of climate change particularly CO₂ emissions on food security. Future research in this area could offer a comprehensive framework more for policymakers and agricultural stakeholders seeking to increase food security, especially with the unpredictable climate conditions.

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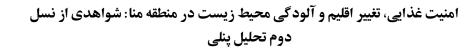
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چکیدہ

امنیت غذایی بدلیل رشد جمعیت، موقعیت جغرافیایی و اقلیمی، یک مساله حیاتی در منطقه خاورمیانه و شمال آفریقا (منطقه منا) است. از دیگر سو بیشتر کشورهای واقع در این منطقه از منابع طبیعی فراوان با محوریت سوختهای فسیلی منفعت میبرند. همچنین مسایل محیط زیستی، بویژه انتشار گازهای گلخانهای ناشی از فعالیتهای تولید و فشارهای ناشی از تغییرات اقلیمی اهمیت امنیت غذایی را برجسته نموده است. در ین مطالعه تاثیر تغییر اقلیم، آلودگیهای محیط زیستی و سایر متغیرها بر امنیت غذایی در منطقه منا طی دوره به ۱۹۹۰ الی ۲۰۱۹ مورد بررسی قرار گرفت. با توجه به وابستگی مقطعی نسل دوم برآوردگر پنلی CS-ARDL مورد استفاده قرار گرفت. نتایج نشان داد مصرف انرژی، سطح اراضی زراعی، انتشار گاز دی اکسید کرین و بارندگی تاثیر مثبت و معنی داری بر امنیت غذایی دارد. بعلاوه شهرنشینی و متوسط دما دارای تاثیر منفی هستند. نتایج آزمون علیت نشان داد که اراضی زراعی و بارندگی دارای رابطه علی یکطرفه با امنت غذایی بوده فه مصرف انرژی، انتشار گاز دی اکسید کرین، شهرنشینی و متوسط دما دارای تاثیر منفی غذایی هستند. نتایج حکی از آن است که ضمن حفظ و افزایش تولید محصولات کشاورزی، باید به ازان غذایی هستند. نتایج میا میان داد که اراضی زراعی و بارندگی دارای رابطه علی یکطرفه با امنت غذایی بوده اقلیمی و تاثیرات محیطی زیستی تولید نیز توجه نمود.