A Panel Data Approach for Investigation of Gross Domestic Product (GDP) and CO₂ Causality Relationship

I. Saleh¹, S. Abedi¹*, and S. Abedi²

ABSTRACT

One of the current challenges and complications in the world is the climate change and global warming, which has numerous and varied effects and consequences in different regions. In this regard, the effects of economic activities on the increase in greenhouse gases and also the effects of greenhouse gases on economic activities have become increasings controversial. In this study, an investigation was done upon the bidirectional causality relationship between real gross domestic product per capita and carbon dioxide emissions per capita in different countries. For this purpose, the Vector Auto-Regression Model with the micro panel application was used and the World Bank member countries were divided into different groups. Results indicated bidirectional causality relationship between Gross domestic product (GDP) and CO₂ for three groups of countries. In addition, there was a one-way causal relationship from GDP to carbon dioxide volume for subgroups of countries with high average economic growth rate (HGR) and the rest of the world countries (ROW). This means that, to accomplish the international goals of decreasing the emissions of pollutant gases, collaboration between HGR and ROW group of the countries with industrial countries is indispensable. Moreover, heterogeneous non-causality test for Iran suggests that the economic activities are having increasingly negative environmental impacts on the country.

Keywords: Carbon dioxide, Causality test, Gross domestic product, Micro panel, Vector auto-Regression.

INTRODUCTION

The Intergovernmental Panel on Climate Change (IPCC) reported that since 1850, the year when systematic temperature-keeping began (IPCC, 2007), eleven of the 12 warmest years occurred during 1995 to 2006. This, along with other evidence of this warming such as the number of shrinking mountain glaciers, thawing permafrost, earlier breakup of rivers and lakes ice, lengthening of mid- to high-latitude growing seasons, shifts of plants, insects, and animal ranges, earlier tree flowering, insect emergences and egg laying in birds and changes in precipitation patterns indicate that the trend of global warming is virtually certain (IPCC, 2007).

With the start of the industrial revolution in the second half of the nineteenth century and the human beings’ ever-increasing need for energy and fossil fuels such as coal, oil, and natural gases, the earth has encountered an increase in polluting gases in the atmosphere (e.g. carbon dioxide and methane). In this period of time, the amount of the present carbon dioxide and methane in the earth’s atmosphere has increased 13% and 151%, respectively (IPCC, 2007).

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Therefore, one of the current challenges and complications in the world is the climate change and global warming, which has numerous and varied effects and consequences in different regions. Since almost 60 percent of the greenhouse effects are related to human activities in CO₂ emissions, this gas is considered the most important greenhouse gas in regard to global warming (IPCC, 2007).

Two different theories have been proposed about the effects of economic activities on the environment. Some researchers believe that since economic growth requires a greater use of raw material and energy, the economic activities will unavoidably cause vast environmental damage (Perman et al., 2003; Acaravci and Ozturk, 2010; Hossain, 2011; Pao and Tsai, 2011; Wang et al., 2011). Some, in contrast, insist that since economic growth brings about an increase in products and services and promotes welfare, it will gradually encourage and improve motivations for supporting the environment and, thus, will be prioritized in the policy programs of countries; therefore, economic growth is a need for qualitative improvement of the environment. The World Bank study (1992) and Everett et al. (2010) are among the studies which are in this group.

On the other hand, some empirical evidence attest to the possibility of the impact of the environmental effects on economic efficiency. Pigou (1920) is one of the pioneers of this theory. Other investigations, which have focused upon the relationship between greenhouse gases and the economic index, are those done by Dinda and Coondoo (2006), Azomahou et al. (2006), Liu (2006), Masih et al. (2010), Maddison and Rehdanz (2008), Ferda (2008), Lee (2009), Pao and Tsai (2010), Moghadas and Golriz (2011), and Hatzigeorgiou et al. (2011) have also focused on the bidirectional causal relationship between gross domestic product and the volume of greenhouse gases or, in particular, carbon dioxide. Therefore, the effect of economic activities on the rise of the volume of greenhouse gases and the effect of greenhouse gases on economic activities has been highly controversial.

After frequent warnings given by environmental specialists about the effect of greenhouse gases on global warming, representatives from more than 160 countries ratified the Kyoto Protocol with the aim of obliging different countries to lower the emissions of these gasses. (It might be noticed that the performance of the protocol was launched in February 2005 after Russia joined.)

On the basis of this protocol, 36 industrial countries made a commitment to lower their greenhouse gases emissions rate by about 5.2 percent from 1990 levels between 2008 and 2012. In this protocol, it has been mentioned that economic growth and alleviation of poverty, both causing an increase in the emissions of greenhouse gases, are the rights of the developing countries.

The countries that signed the protocol pledged to pay the pollution costs to the countries whose greenhouse gas emissions were lower than the permissible level, if the emissions rate of their greenhouse gases, which affects global warming, did not change or increase. However, the US, China, and India did not sign the Kyoto Protocol. The United States believed that lowering the emissions of greenhouse gases, which is done by decreasing industrial activities, would damage the economy of the country. The USA asserts that until countries like China and India, which have experienced rapid economic growth and also produced large amounts of greenhouse gasses, do not sign this protocol, considerable changes in the atmospheric state cannot be anticipated. Hence, it seems that in response to the issues raised, it is necessary to study the economy of countries which are the main sources of emissions of greenhouse gasses as well as the economic communities that are more vulnerable to the rise of greenhouse gases.

Iran is amongst the second group of the countries committed to the Kyoto Protocol, which are mostly developing countries.
According to unpublished information of Iran’s Ministry of Energy, more than 75 percent of the emissions of greenhouse gases in Iran is because of high consumption of fuel and cheap energy.

This study tested the hypothesis of the existence of a bidirectional causal relationship between gross domestic product and emissions of carbon dioxide for different countries, with a focus on Iran. In addition, by studying the mentioned concept for three groups of countries, namely, Organization for Economic Co-operation and Development member countries, countries with high rate of average economic growth, and other countries, this study aimed to find out the economy of which group of countries is the main factor in the emissions of carbon dioxide and its consequences and which group is more affected by this phenomenon. The innovation of the current study is the application of the data set of a greater number of countries and the investigation of countries in three main economic groups.

**MATERIALS AND METHODS**

In the literature, Granger (1969) causality test, which uses the panel data approach, includes two main groups of methods. The first group considers estimation and testing vector autoregression (VAR) coefficient using autoregressive coefficients and regression coefficient slopes as variables in panel data. This approach was applied by Holtz-Eakin et al. (1988), Nair-Reichart and Weinhold (2001), Choe (2003), Costantini and Martini (2009), Lean and Smyth (2009), Sinha (2009) and Lau et al. (2011). The second group suggests autoregressive coefficients and regression coefficients slopes as constant values. This method was developed by Hurlin and Venet (2001), Hurlin (2004a, b), Hansen and Rand (2004), Amiri and Gerdtham (2013). The second method is employed in this study, due to its compatibility with our data sets with relatively short time periods compared to large numbers of cross-section units. Following Hurlin and Venet (2001), two covariance stationary variables, marked by $x$ and $y$ were considered for $T$ periods and $N$ cross-section units of observations.

According to Granger (1969) causality procedure, for each individual $i \in \{1, N\}$, $x_{i,t}$ explains $y_{i,t}$, if the predictive power of $y_{i,t}$ employing the data of all variables including $x_{i,t}$, is greater than when excluding it. Since it is not practically possible to completely apply the optimum predictors, this study considers only the linear ones. Therefore, a VAR model which is to be used for panel data is presented below. For each cross-section unit $i$ and period $t$, the following model is estimated:

$$y_{i,t} = \sum_{k=1}^{p} \beta_k y_{i,t-k} + \sum_{k=0}^{p} \theta_k x_{i,t-k} + u_{i,t} \quad (1)$$

Where, $u_{i,t} = \alpha + \varepsilon_{i,t}$ and $\varepsilon_{i,t}$ are i.i.d and $N(0, \sigma^2)$. It is assumed that autoregressive coefficients $\beta_k$ and the regression coefficient $\theta_k$’s are constant for $k \in [1, N]$. It is further assumed that parameters $\beta_k$ are identical for all individual, while the coefficient $\theta_k$ could have an individual dimension. In other words, the model employed in this study was panel data model with constant coefficients. Finally, the residuals are assumed to satisfy the standard properties.

There are a lot of advantages in using panel data models. In addition to providing a large number of observations, increasing the degrees of freedom and reducing the co-linearity among explanatory variables, it obviously improves the efficiency of Granger causality tests (Hurlin and Venet, 2001). In testing causality with panel data, the possibility of heterogeneity among cross-section units should be considered. Inmate cross sectional disparities is the first source of heterogeneity. A pooled estimation without heterogeneous intercepts results in a
bias of the slope estimates and may lead to fallacious inference of causality tests (Hurlin, 2004a). Heterogeneous regression coefficients $\theta_k$ are the second basis of heterogeneity and cause more difficulties than the first reason. Furthermore, different heterogeneity sources of the data-generating process should be considered in the causality test using panel data sets. Hence, different kinds of causality hypotheses are to be tested in a panel data set framework. The first test procedure, named as homogenous and instantaneous non-causality hypothesis (HINC), is directed towards testing whether or not the $\theta_k$’s of $x_{i,t-k}$ are simultaneously null for all individuals $i$ and all lags $k$. The hypotheses are:

$$H_0 : \theta_k = 0 \forall i \in [1, N] \forall k \in [0, p]$$

$$H_1 : \theta_k \neq 0 \exists (i,k)$$

For testing $Np$ linear restrictions (2), the following Wald statistic is calculated:

$$F_{HINC} = \frac{(SSR_r - SSR_o)/(Np)}{SSR_o / [NT - N(1 + p) - p]}$$

(3)

Where, $SSR_o$ stands for the sum of squared residuals for model in (1) and $SSR_r$ for the restricted sum of squared residuals under $H_o$. If individual effects, $\alpha$, are assumed to be fixed, $SSR_o$ and $SSR_r$ are $SSR$ obtained from the maximum likelihood (ML) estimation that corresponds in this case to the fixed effects (FE) estimator.

If the HINC hypothesis is rejected, there are two possibilities. The first one is the homogenous causality hypothesis (HC) and takes place if all the coefficients $\theta_k$ are identical for all lags $k$ and are statistically different from zero. In other words, we are testing whether or not $\theta_k$’s in (1) are equal to each other. Therefore, the following hypotheses are tested:

$$H_0 : \theta_k^i = \theta_k^j \forall i, j \in [1, N] \forall k \in [0, p]$$

$$H_1 : \theta_k^i \neq \theta_k^j \exists (i, j, k)$$

(4)

In order to test (4), the following statistic is calculated:

$$F_{HC} = \frac{(SSR_r' - SSR_o')/[p(N-1)]}{SSR_o' / [NT - N(1 + p) - p]}$$

(5)

Where, $SSR'_r$ is the restricted sum of squared residuals under $H_o$. As in the case of HINC, if individual effects, $\alpha$, are assumed to be fixed, the ML estimator is consistent with the FE estimator.

If the HC hypothesis is also rejected, the process is non-homogenous and no homogeneous causality relationship can be obtained. Nonetheless, it does not entail rejection of any causality relationships between two variables and for a sub-group of cross-section units it may still be possible that there exist causality relationships and the variable $x$ causes the variables $y$. The last step, therefore, is to test heterogeneous non-causality hypothesis (HENC). The hypotheses are:

$$H_0 : \theta_k^i = 0 \forall i \in [1, N] \forall k \in [0, p]$$

$$H_1 : \theta_k^i \neq 0 \forall i \in [1, N] \forall k \in [0, p]$$

(6)

Under this case, the nullity of all the coefficients of the lagged explanatory variable $x_{i,t-k}$ is tested for each cross-section unit. For testing (6), the following statistic is calculated:

$$F_{HENC} = \frac{(SSR_r'' - SSR_o'')/p}{SSR_o'' / [NT - N(1 + 2p) + p]}$$

(7)

Where, $SSR''_r$ is sum of squared residuals found in (1), when the nullity of the $k$ coefficients associated with the variable $x_{i,t-k}$ are imposed only for the cross-section unit $i$. These $N$ individual tests identify the cross-section units for which there are no causality relationships. If the HENC hypothesis is not rejected, the variable $x$ does not cause the variable $y$ for a sub-group of cross-section units. The causality relationship is relevant only for a sub-group of cross-section units. This hypothesis can
be considered as the consequence of data-generating process heterogeneity.

The following two models are estimated for each of the three groups:

\[
GDPP_{t,i} = \sum_{k=0}^{p} \beta_k GDPP_{t-i-k} + \sum_{k=0}^{p} \theta_k CP_{t-i-k} + u_{t,i},
\]

(8)

\[
CP_{t,i} = \sum_{k=0}^{p} \beta_k CP_{t-i-k} + \sum_{k=0}^{p} \theta_k GDPP_{t-i-k} + u_{t,i},
\]

(9)

Where, GDPP is real gross domestic product per capita and CP is CO\textsubscript{2} emissions per capita. Before estimating Equations (8) and (9), the optimum lag length should be determined for both variables in each country group using Akaike Information Criterion (AIC). Following the adoption of a model with white noise residuals, both homogenous and instantaneous non-causality (HINC) and homogenous causality (HC) hypotheses were tested using (8) and (9). Later, heterogeneous non-causality hypothesis (HENC) was tested for Iran in the other countries group.

Countries were classified into three consistent groups in order to analyze the information of countries homogenously and present appropriate answers to the questions of the study. The questions are: in which group of countries does CO\textsubscript{2} emissions, as the most important greenhouse gas, affect economic growth more, and in which group does it have less impact? On the other hand, in which type of countries, does economic growth affect CO\textsubscript{2} emissions? After eliminating the countries with missing observations, we formed three groups of countries. The first group included thirty countries which were members of OECD. The second group (HGR) was made up of seventeen countries which had a GDP growth rate of more than 4 percent in our study period according to the World Bank’s World Development Indicators. The third group (ROW) contained fifty-one countries which were included in neither of the two other groups. All the data were converted into natural logarithms prior to conducting the analysis. The data of this study, which provided for testing the bidirectional causality between GDP and CO\textsubscript{2} in a panel data setting, were derived from the World Bank’s data base between the years of 1990-2004.

RESULTS AND DISCUSSION

The first step in exploring the bidirectional causality between CO\textsubscript{2} emissions per capita and gross domestic production is to test whether the variables at hand contain unit roots. In this study, we used Levin, Lin and Chu (2002) and PP panel unit root tests proposed by Maddala and Wu (1999). The null hypothesis in these panel unit root tests is that the panel series has a unit root (non-stationary). The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1. The results of the panel unit root tests for GDP and CO\textsubscript{2} are shown in Table 1.

<table>
<thead>
<tr>
<th>Country Group</th>
<th>Variable</th>
<th>Levin, Lin and Chu</th>
<th>PP</th>
<th>Levin, Lin and Chu</th>
<th>PP</th>
</tr>
</thead>
<tbody>
<tr>
<td>HGR\textsuperscript{a}</td>
<td>GDP\textsuperscript{d}</td>
<td>-1.90\textsuperscript{**}</td>
<td>47.7\textsuperscript{***}</td>
<td>-2.06\textsuperscript{**}</td>
<td>72.51\textsuperscript{***}</td>
</tr>
<tr>
<td></td>
<td>CO\textsubscript{2}</td>
<td>1.96\textsuperscript{**}</td>
<td>37.04\textsuperscript{***}</td>
<td>-1.39\textsuperscript{***}</td>
<td>12.59</td>
</tr>
<tr>
<td>OECD\textsuperscript{b}</td>
<td>GDP</td>
<td>-4.30\textsuperscript{*}</td>
<td>114.71\textsuperscript{*}</td>
<td>6.15</td>
<td>110.25\textsuperscript{*}</td>
</tr>
<tr>
<td></td>
<td>CO\textsubscript{2}</td>
<td>-1.29\textsuperscript{***}</td>
<td>85.02\textsuperscript{*}</td>
<td>5.32</td>
<td>130.71</td>
</tr>
<tr>
<td>ROW\textsuperscript{c}</td>
<td>GDP</td>
<td>-5.03\textsuperscript{*}</td>
<td>170.71\textsuperscript{*}</td>
<td>5.83</td>
<td>263.51</td>
</tr>
<tr>
<td></td>
<td>CO\textsubscript{2}</td>
<td>-5.98\textsuperscript{*}</td>
<td>121.71\textsuperscript{*}</td>
<td>3.16</td>
<td>140.61</td>
</tr>
</tbody>
</table>

\textsuperscript{a} High Growth Rate, \textsuperscript{b} OECD: Organisation for Economic Co-operation and Development, \textsuperscript{c} Rest of Other Word, \textsuperscript{d} Gross Domestic Product, \textsuperscript{**} and \textsuperscript{***}: Denote statistical significance at the (1%), (5%) and (10%) level, respectively.
the log levels of CO2 and GDP are statistically significant for all country groups with the exception of the PP test applied to CO2 with existence of trend for HGR country. However, because of the non-existence of a trend for this variable, taken as a whole, the log levels results suggest that both variables are panel stationary in general.

The next step in an attempt to search for causality is to choose the lag lengths for both variables. Table 2 presents AIC figures for each country group. Consequently, for High Growth Rate countries (HGR) and Organization for Economic Co-operation and Development members (OECD), we chose three lags for variables GDP and CO2. In the case of Rest-Of-the-World group (ROW), the corresponding lag lengths were four.

Before showing estimation and statistical inference of regression, it was necessary to determine the method of estimation of consolidated data. In order to choose between methods of common effects (mix regression model) and fixed effects (panel data), an F-Limer test was used. As can be seen in the results of Table 3, the compilation data estimation method (common effect method) is rejected. F-Limer test results for all country groups showed that the common effect method for estimating the regression models was not suitable. In other words, the intercept was different for different units or otherwise, there was individual or group effects and should be used fixed effects methods to estimate the models. In the next step, in order to show that model should be estimated by a fixed effects method or a random effects method, the Hausman test was performed. The results of the test are presented in Table 3.

In Table 3, the P-value of the Hausman statistics also show that fixed effects method used to estimate the model is a more appropriate option. Therefore, according to the results of the F-Limer test and the Hausman test, the most appropriate method to estimate parameters and test hypotheses is

Table 2. Optimum lag length using Akaike Information Criterion. 

<table>
<thead>
<tr>
<th>Country Group</th>
<th>LAG 1</th>
<th>LAG 2</th>
<th>LAG 3</th>
<th>LAG 4</th>
<th>LAG 5</th>
</tr>
</thead>
<tbody>
<tr>
<td>HGR</td>
<td>2.36</td>
<td>2.28</td>
<td>2.16</td>
<td>2.19</td>
<td>-</td>
</tr>
<tr>
<td>OECD</td>
<td>0.69</td>
<td>0.62</td>
<td>0.46</td>
<td>0.50</td>
<td>-</td>
</tr>
<tr>
<td>ROW</td>
<td>1.21</td>
<td>1.19</td>
<td>1.14</td>
<td>1.13</td>
<td>1.14</td>
</tr>
</tbody>
</table>

* Indicates lag order selected by the criterion. Source: Research findings.  

Table 3. Results of F-Limer Test and Hausman Test.

<table>
<thead>
<tr>
<th>Test type</th>
<th>Country group</th>
<th>Type of Relationship</th>
<th>Value test result</th>
<th>Probability</th>
</tr>
</thead>
<tbody>
<tr>
<td>F-Limer</td>
<td>HGR</td>
<td>GDP → CO2</td>
<td>1.61</td>
<td>0.0699</td>
</tr>
<tr>
<td></td>
<td></td>
<td>CO2 → GDP</td>
<td>2.23</td>
<td>0.0058</td>
</tr>
<tr>
<td></td>
<td></td>
<td>GDP → CO2</td>
<td>1.52</td>
<td>0.0437</td>
</tr>
<tr>
<td></td>
<td>OECD</td>
<td>GDP → CO2</td>
<td>2.12</td>
<td>0.0009</td>
</tr>
<tr>
<td></td>
<td>ROW</td>
<td>GDP → CO2</td>
<td>2.20</td>
<td>0.0000</td>
</tr>
<tr>
<td></td>
<td></td>
<td>CO2 → GDP</td>
<td>3.38</td>
<td>0.0000</td>
</tr>
<tr>
<td></td>
<td></td>
<td>GDP → CO2</td>
<td>18.42</td>
<td>0.0051</td>
</tr>
<tr>
<td></td>
<td>HGR</td>
<td>GDP → CO2</td>
<td>35.49</td>
<td>0.0000</td>
</tr>
<tr>
<td></td>
<td>OECD</td>
<td>GDP → CO2</td>
<td>41.85</td>
<td>0.0000</td>
</tr>
<tr>
<td></td>
<td>ROW</td>
<td>GDP → CO2</td>
<td>59.43</td>
<td>0.0000</td>
</tr>
<tr>
<td></td>
<td></td>
<td>CO2 → GDP</td>
<td>104.48</td>
<td>0.0000</td>
</tr>
</tbody>
</table>

|                      | GDP → CO2     | 160.67             | 0.0000             |

Source: Research findings.  

[952]
fixed effects method.

In the next step, Equations (8) and (9) were estimated for each country group in order to test HINC and HC hypotheses. Table 4 demonstrates the results of testing these two types of homogenous causality hypotheses. The results show that homogenous and instantaneous non-causality are rejected at one percent levels of significance for GDP-to-CO2 and CO2-to-GDP causality in all country groups. Rejecting the null hypothesis of HINC means that there exists a causality relationship between GDP and CO2 in these groups. The next step is whether the causality is an overall causality for each country group or sourced from causality relations for individual countries. On the basis of homogeneous causality test results, the existence of HC is rejected. It means that there is a heterogeneous one-way causality relationship from GDP to CO2 for sub-groups countries with HGR and also for sub-groups of ROW members at 1% and 10% levels of significance, respectively.

A number of studies have been carried out in order to investigate the causality relationship between CO2 emissions per capita and GDP per capita for sub-groups of countries such as Taskin and Zaim (2000), and Dijkgraaf and Vollebergh (2001). These studies showed an inverse-U-shape relationship between the mentioned variables in their data set. However, studies by Liu (2006), Masih et al. (2010), Maddison and Rehdanz (2008), Ferda (2008), and Lee (2009), depending upon the country (countries) they analyzed, indicated different interpretations of the relationship between CO2 emissions and GDP value.

The final step was to discover the existence of causality in Iran. Since the homogenous hypotheses for GDP-to-CO2 relationship were rejected for the ROW group including Iran, the heterogeneous non-causality test was carried out. Results of HENC hypotheses for GDP-to-CO2 for Iran are presented at Table 5.

At the 5% significance level, the results show that, for Iran, the null hypothesis indicating that the GDP does not cause CO2, is rejected. This evidence indicates that there is a causal relationship running from GDP to CO2. It means that GDP strongly (Granger) causes CO2, which implies that CO2 is so pervasive in the economy that the economic growth actually increases the CO2 emissions over time.

Table 6 shows the regression results for the Environmental Kuznets Curve for Iran. We clearly reject the existence of an inverted “U” relationship, as long as, the estimated t ratios of the CO2 quadratic term is not significant at any significance level tested. Thus, according to the results, the typical

Table 4. Test results for causality hypotheses.

<table>
<thead>
<tr>
<th>Country group</th>
<th>Test</th>
<th>From GDP to CO2</th>
<th>From CO2 to GDP</th>
</tr>
</thead>
<tbody>
<tr>
<td>HGR</td>
<td>HINC&lt;sup&gt;d&lt;/sup&gt;</td>
<td>9.60&lt;sup&gt;*&lt;/sup&gt;</td>
<td>25&lt;sup&gt;*&lt;/sup&gt;</td>
</tr>
<tr>
<td></td>
<td>HC&lt;sup&gt;e&lt;/sup&gt;</td>
<td>6.12&lt;sup&gt;*&lt;/sup&gt;</td>
<td>1.06</td>
</tr>
<tr>
<td>OECD&lt;sup&gt;b&lt;/sup&gt;</td>
<td>HINC</td>
<td>6.15&lt;sup&gt;*&lt;/sup&gt;</td>
<td>6.22&lt;sup&gt;*&lt;/sup&gt;</td>
</tr>
<tr>
<td></td>
<td>HC</td>
<td>1.13</td>
<td>1.001</td>
</tr>
<tr>
<td>ROW&lt;sup&gt;c&lt;/sup&gt;</td>
<td>HINC</td>
<td>6.39&lt;sup&gt;***&lt;/sup&gt;</td>
<td>4.14&lt;sup&gt;*&lt;/sup&gt;</td>
</tr>
<tr>
<td></td>
<td>HC</td>
<td>1.39&lt;sup&gt;***&lt;/sup&gt;</td>
<td>1.14</td>
</tr>
</tbody>
</table>

Source: Research findings. <sup>a</sup> High Growth Rate, <sup>b</sup> Organisation for Economic Co-operation and Development, <sup>c</sup> Rest of Other Word, <sup>d</sup> Homogenous and Instantaneous Non-Causality, <sup>e</sup> Homogenous Causality. *,**, and ***: Reject H<sub>0</sub> at 1%, 5%, and 10% levels of significance, respectively.

Table 5. Results of HENC hypotheses for Iran.

<table>
<thead>
<tr>
<th>Country group</th>
<th>Test</th>
<th>From GDP to CO2</th>
<th>Probability</th>
</tr>
</thead>
<tbody>
<tr>
<td>Iran</td>
<td>HENC&lt;sup&gt;**&lt;/sup&gt;</td>
<td>5.322297</td>
<td>0.0298</td>
</tr>
</tbody>
</table>

Reference: Research findings, <sup>**</sup> Heterogeneous Noncausality
Table 6. Environmental Kuznets Curve for Iran.

<table>
<thead>
<tr>
<th>Variable</th>
<th>Coefficient</th>
<th>T-ratio</th>
</tr>
</thead>
<tbody>
<tr>
<td>constant</td>
<td>10.389</td>
<td>0.56</td>
</tr>
<tr>
<td>$\text{CO}_2$</td>
<td>64.53</td>
<td>1.08</td>
</tr>
<tr>
<td>$\text{CO}_2^2$</td>
<td>23.34</td>
<td>1.45</td>
</tr>
</tbody>
</table>

Source: Research findings.

The inverted “U” shape of the EKC is not confirmed yet with our data set for Iran. The EKC model study also found that squared $\text{CO}_2$ did not contribute significantly to explain changes of GDP. A more appropriate interpretation of the model is that, perhaps, there were some important variables that were not incorporated in the model.

CONCLUSIONS

In order to answer the research question of this study concerning existence of a bidirectional relationship between real GDP per capita and $\text{CO}_2$ emissions per capita, we divided the countries of our data set into three groups according to their economic diversities. Results showed that gross domestic product per capita for three groups of ROW, OECD and HGR countries affects their $\text{CO}_2$ emissions per capita. This means that cooperation of these countries along with the industrial countries in order to achieve international goals of decreasing the pollutant gases is indispensable. In order for these countries to obtain an economic growth rate higher than that of the environmental polluting gases, technology would perform a vital role. Promoting technology and cooperation among countries, especially between the two groups of industrial countries and developing countries, might be a proper response to the challenge of greenhouse gases all around the world. The government and the private sector need to invest in appropriate projects and cooperate as well. As a matter of fact, governments are able to prevent the increasing trend of climatic changes by adopting appropriate policies and supporting the private investors on the basis of a market economy.

On the other hand, for ROW, OECD and HGR, the results revealed that their economy is under the influence of greenhouse gas emissions. Therefore, it seems appropriate that these countries encourage transferring clean technologies to other countries by accelerating the flow of international aids. It is only possible in the long term to develop cleaner and cheaper energy resources in order to guarantee permanent economic growth. Therefore, developing appropriate policies on climate change problems might be an inseparable factor of sustainable development in the near future. Some solutions recommended recently for decreasing greenhouse gases include using solar, wind, and nuclear energy, as well as fossil fuel consumption management and silviculture programs.

In the case of Iran, according to our study, gross domestic product affects $\text{CO}_2$ emissions. It indicates that the economic and human activities are having increasingly negative environmental impacts on the country. Therefore, it seems necessary to implement appropriate policies in order to decrease the destructive effects of climatic change. Besides, with the presence of global markets, it is possible to utilize the available opportunities of new knowledge and technologies in order to decrease greenhouse gases, increase efficiency, improve energy resource productivity, manage consumption properly, and also promote clean and renewable energy resources.

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رویکرد داده‌های ترکیبی برای بررسی تاثیر تولید ناخالص داخلی و دی اکسید کربن

1. صالح س. عابدی و س. عابدی

چکیده

یکی از مشکلات و مشکلات فعّال در جامعه جهانی، بروز تغییرات آب و هوایی و گرمایش جهانی است که اثرات و تبعات فراوانی و متغیری در نقاط مختلف دنیا دارد. در این رابطه، از سویی اثر تغییرات دانسته‌های اقتصادی بر افزایش حجم گازهای گلخانه‌ای و از سویی دیگر، تأثیر منازع گازهای گلخانه‌ای بر تغییرات اقتصادی به طور قابل‌توجهی متحال می‌شود. این مطالعه بررسی فرآیند ارتباط عضویت دانسته‌های گازهای گلخانه‌ای و دانسته‌های ناخالص داخلی دارای است. برای این منظور، از مدل بردار خودکارگری با کاربرد ماکروپولی استفاده شده و گزارشات عضویت گازهای گلخانه‌ای مختلف تخمین شده‌اند. نتایج حاکی از وجود ارتباط دوکمی دانسته‌های هم‌زمان برای سه گروه کشور مورد مطالعه می‌باشد. این نتایج به‌علاوه نتایج حاکی از تأثیر منازع حجم گازهای دی اکسید کربن برای زیرمجموعه از کشورهای با رشد متسویم اقتصادی بالا (HGR) و سایر کشورها (ROW) به چشم می‌خورد. این بدان معناست که باید تلیه به اهداف بین‌المللی کاهش انتشار گازهای دی اکسید کربن با روش‌های مختلفی مورد نظر باشد. در کار توجه‌های مختلفی به پیشنهادات کشورهای ROW و HGR مراجعه شده است که می‌تواند به سیاست‌های کشورهای ROW مرتبط با این موضوع کمک کند.

علی‌رغم بررسی آزمون‌های نامناسب برای ایران، نتایج آزمون مشابه نامناسب برای ایران بایانگر آن است که فعالیت‌های اقتصادی دارای اثرات

زیست محیطی متفاوت قابل توجهی بر کشور است.