

# Estimating Daughter Yield Deviation and Validation of Genetic Trend for Somatic Cell Score in Holstein Cattle Using Random Regression Test Day Model

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## ABSTRACT

The objective of this study was to estimate Daughter Yield Deviations (DYDs) of bulls and Yield Deviations (YDs) for cows using a random regression model and validation of genetic trend using estimated DYDs and Method II of Interbull for test-day records of Somatic Cell Score (SCS) in the first lactation of Iranian Holsteins. Data set included the 108995 test day records collected by the Animal Breeding Center of Iran from 2001 to 2010. Results of the present study indicated that variation in YDs of cows at different stages of lactation corresponds closely with their Estimated Breeding Values (EBVs). Because YDs and DYDs are considered as an additional measure of an animal's genetic merit, their correlation with EBVs is very important. The correlation between DYDs and EBVs of bulls for SCS was 0.88. High correlation estimates between DYDs and EBVs indicated that, in addition to EBV, the DYD can be an appropriate measure for dairy cattle breeding programs. The correlation increased with increase in the number of bull daughters and the average number of test-days of daughters. Estimated DYDs for each production year were used to validate the genetic trend obtained from the model which was used for genetic evaluation. Results indicated that genetic trend for SCS in the first lactation of Iranian Holsteins was slightly overestimated.

**Keywords:** Dairy cow, Estimated breeding value, Genetic progress, Mastitis, Validation of genetic trend.

## INTRODUCTION

Mastitis, or inflammation of the mammary gland, is one of the most complex and costly diseases affecting dairy cattle. Costs due to clinical mastitis include lower milk production, poor milk quality, discarded milk, veterinary costs, and premature culling of cows (Kadarmideen and Pryce, 2001; Koivula *et al.*, 2004). Selection against mastitis, in countries where incidences are not recorded, is carried out indirectly by selecting against Somatic Cell Count (SCC) (Mrode and Swanson, 1996). Daily Somatic Cell Score (SCS) has usually been analyzed as repeated measurements of the same trait. However, the

genetic correlations between SCS at different stages of lactation are less than unity, which violates the assumptions of the repeatability model (Reents *et al.*, 1994; Mrode *et al.*, 1996; Haile Mariam *et al.*, 2001; Ødegard *et al.*, 2003). The recent trend in dairy cattle genetic evaluations is towards the application of Random Regression Models (RRM) using Test Day (TD) records (Mrode and Swanson, 2004) that eliminated the deficiency of repeatability model.

Besides the Estimated Breeding Value (EBV), Yield Deviation (YD) of cows and Daughter Yield Deviation (DYD) of bulls are important quantities used in dairy cattle selection (Szyda *et al.*, 2008). The YD is a weighted average of the cows' yields adjusted

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for all effects of the model other than genetic merit and error. The *DYD* of bulls is the average performance of their daughters that are adjusted for fixed and non-genetic random effects of the daughters and genetic effect of their mates (VanRaden and Wiggans, 1991; Liu *et al.*, 2004; Freyer *et al.*, 2002). *DYD* is not regressed on breeding value of bulls and is the most independent and accurate measure of phenotypic performance of a bull's daughters (Van Raden and Wiggans, 1991; Liu *et al.*, 2004). Van Raden and Wiggans (1991) showed the calculation method of *YD* and *DYD* for repeatability animal model. Mrode and Swanson (2002) presented this calculation for a random regression model. Liu *et al.* (2003) developed a method for calculation of *DYD* under general multiple trait models.

For SCS records, Mrode and Swanson (2004) reported *DYD* for Holstein-Friesian heifers. Also Liu *et al.* (2004) calculated *DYD* for SCS in Holstein, Red and Jersey dairy cattle from Austria, Germany and Luxembourg. Calculation of *YD* and *DYD* in Iranian Holsteins was performed by Sheikhlou *et al.* (2009) for milk and fat traits using repeatability animal model. Khazadeh *et al.* (2013) calculated *YD* and *DYD* for production traits of Iranian Holsteins using both repeatability animal and Random Regression Test Day Models (RRTDMs). However, the *YD* and *DYD* have not been calculated for SCS records until now in Iranian dairy cows. Hence, the objective of this study was to estimate *YD* and *DYD* for SCS using RRM and its application for the genetic evaluation of Iranian Holsteins.

## MATERIALS AND METHODS

### Data and Model

A total of 108,995 daily SCS records from 2001 to 2010 for the first lactation of the Iranian Holsteins were obtained from the Animal Breeding Center of Iran and analyzed by the following RRM using the AIREML algorithm of the WOMBAT program (Meyer, 2006):

$$Y_{ismptv} = YS_s + HTD_m + \sum_{f=1}^2 C_f (age_n)^f + \sum_{r=0}^k \beta_r \phi_r (dim_t) + \sum_{r=0}^{k_p-1} \alpha_{pr} \phi_r (dim_t) + \sum_{r=0}^{k_{pr}-1} \gamma_{pr} \phi_r (dim_t) + e_{imptv}$$

Where,  $Y_{imptv}$  is test day SCS record  $i$  obtained at  $DIM_t$  of cow  $p$  calved at the  $n^{th}$  age in year-season of calving  $s$  and herd-test date  $m$ ;  $YS_s$  is fixed effect of the  $s^{th}$  year-season of calving;  $HTD_m$  is fixed effect of the  $m^{th}$  herd-test date;  $C_f$  is the  $f^{th}$  fixed regression coefficient for calving age;  $age_n$  is the  $n^{th}$  calving age;  $k$  is the order of fit for fixed regression coefficients ( $k= 3$  or  $4$ );  $\beta$  is the  $r^{th}$  fixed regression coefficient;  $k_a$  is the order of fit for additive genetic random regression coefficients;  $k_p$  is the order of fit for permanent environmental random regression coefficients;  $\alpha_{pr}$  is the  $r^{th}$  random regression coefficient of additive genetic value of  $p^{th}$  cow;  $\gamma_{pr}$  is the  $r^{th}$  random regression coefficient of permanent environmental effect of  $p^{th}$  cow;  $\phi_r (dim_t)$  is the  $r^{th}$  coefficient of Legendre polynomials evaluated at days in milk  $t$ ; and  $e_{imptv}$  is the random residual error.

In general, 16 different models were fitted for the analysis of the data set (Table 1). These models differed in terms of the Legendre polynomials used to fit the covariance functions for additive genetic and permanent environmental effects and in the number of classes for the residual variances. Selection of models was based on Akaike's Information Criterion (AIC) (Akaike, 1973). Model 15 had the lowest AIC value, therefore, it was chosen for the analysis of SCS records.

### Calculating Yield Deviations

Equations to calculate the contribution of information from different sources of random regression coefficients in the random regression model for any animal is

**Table 1.** Different orders of fit for random regression coefficients in this study.

Model	Fixed regression order of fit	$ka^a$	$kp^a$	Residual variance class	$Np^b$	$\text{Log } l^c$	$AIC^d$
1	3	3	3	4	16	-101665.74	203363.48
2	3	3	3	10	22	-101659.68	203363.36
3	3	3	4	4	20	-101582.75	203205.50
4	3	3	4	10	26	-101575.09	203202.18
5	3	4	4	4	24	-101581.36	203210.72
6	3	4	4	10	30	-101573.64	203207.28
7	3	4	5	4	29	-101541.436	203140.87
8	3	4	5	10	35	-101544.567	203159.13
9	4	3	3	4	16	-101583.32	203198.64
10	4	3	3	10	22	-101577.75	203199.50
11	4	3	4	4	20	-101501.88	203043.76
12	4	3	4	10	26	-101495.36	203042.72
13	4	4	4	4	24	-101500.81	203049.62
14	4	4	4	10	30	-101494.24	203048.48
<b>15</b>	<b>4</b>	<b>4</b>	<b>5</b>	<b>4</b>	<b>29</b>	<b>-101466.114</b>	<b>202990.23</b>
16	4	4	5	10	35	-101462.042	202994.08

<sup>a</sup> Orders of fit for additive genetic and permanent environmental effects, respectively; <sup>b</sup> Number of parameters for estimated variance function; <sup>c</sup> Maximum log likelihood, <sup>d</sup> Akaike's Information Criterion; Bold values correspond to the best model.

presented by Mrode and Swanson (2004). The equation for  $YD$  calculation in RRM can be written as follows:

$$YD = (Q'R^{-1}Q)^{-1}(Q'R^{-1}Y_C)$$

Where,  $YD$  is a vector of weighted regressions of the animal's  $TD$  yields adjusted for all effects other than additive genetic effect on orthogonal polynomials for  $DIM$ ,  $Q$  is a matrix of orthogonal polynomials of days in milk of order 4 for random animal effect,  $R$  is a diagonal matrix for residual variances and  $Y_C$  is a vector of test day records of cows that is adjusted for all effects in the model, except the additive genetic and residual effects.  $YD$  was estimated for 12,142 Holstein cows.

### Calculating Daughter Yield Deviations

To calculate  $DYD$  of a bull, only records of his own daughters should be considered, male progeny must be excluded, because they do not have own performance records for production traits (Liu *et al.*, 2003). In

this study,  $DYD$  was calculated as follows (Mrode and Swanson, 2004):

$$DYD = \frac{\sum G^{-1}w_{2prog}q_{prog}(2YD_{prog} - u_{mate})}{G^{-1}w_{2prog}q_{prog}}$$

Where,  $DYD$  is a vector of  $DYD$  of bulls expressed as random regression coefficients;  $G$  is genetic covariance matrix;  $q_{prog}$  equals 1 if other parent of the progeny is known and  $\frac{2}{3}$  if unknown;  $YD_{prog}$  and  $u_{mate}$  are estimated

$YD$  for daughters of bull and breeding value for mates of bull, respectively; and

$$W_{2prog} = (G^{-1}\alpha_{anim} + DIAG)^{-1}DIAG,$$

where,  $\alpha_{par} = 1, 2/3$  or  $1/2$  if both, one, or neither parents are known, respectively, and  $\alpha_{prog} = 1$  if animal's mate is known and  $2/3$  if unknown. Note that  $\alpha_{anim} = 2\alpha_{par} + 0.5\alpha_{prog}$

and  $DIAG = Q'R^{-1}Q$ . Computation of  $DYD$  was performed using the computing strategy of  $DYD$  illustrated by Mrode and Swanson (2004). In the present study, bulls with number of daughters less than 10 were removed from the data set.



### Validation of Genetic Trend

Boichard *et al.* (1995) described a method to validate the estimation of genetic trend using *DYD* (Method II of Interbull). Genetic trend validation comprises the estimation of a regression coefficient of *DYD* on the production year (Szyda *et al.*, 2008). After estimating *DYD*, regression coefficient of *DYDs* on the bulls' birth years was estimated using the regression procedure of the SAS software package (SAS, 2002). The model is validated by Interbull when the absolute value of the regression coefficient is less than  $0.01 \times SD$ , where *SD* is the genetic Standard Deviation for the trait (van Steenbergen *et al.*, 2005).

### RESULTS AND DISCUSSION

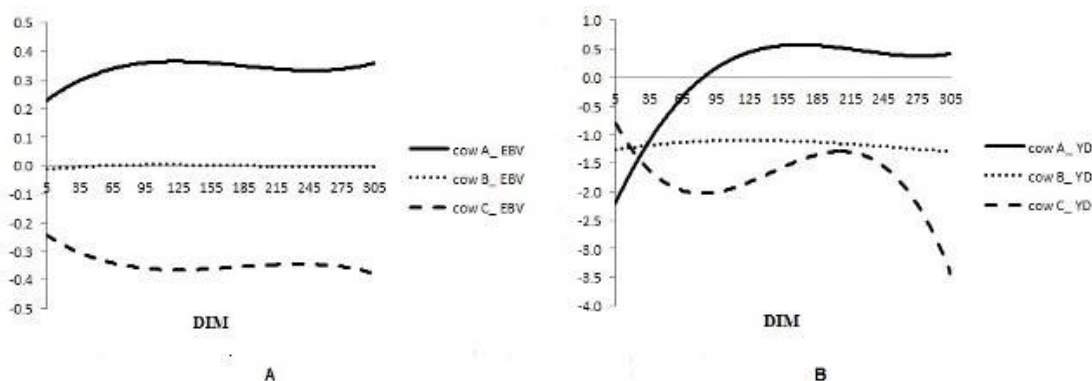
Simple descriptive statistics calculated for *YDs* and *EBVs* of cows for 305 days are presented in Table 2. Because *YD* and *DYD* in RRM expresses in the form of regression coefficients, any linear function of the regression coefficient estimates can be derived for individual cows and bulls,

respectively. In the present study, *EBVs* and *YDs* were calculated for individual days of lactation for three cows (with positive *EBV*, with negative *EBV*, and with *EBV* close to zero). These *EBVs* and *YDs* are presented in Figure 1. As shown in Figure 1, in these cows, the trend in daily *YDs* at different stages of lactation corresponds closely to their *EBVs*. As indicated in Table 2, variation of *YDs* is greater than the variation of *EBVs* which is sensible for the three cows presented in Figure 1. The *DYD* statistics for SCS records and correlation of *DYD* with *EBV* by number of daughters are presented in Table 3. As average number of daughters increased, correlation between *DYD* and *EBV* increased.

As shown in Table 4, correlation between *EBVs* and *DYDs* are much more variable when the average number of test-days per bull's daughters is considered. As average number of test-days increased, correlation also increased. High correlations ( $> 0.90$ ) between *DYD* and *EBV* were observed, on average, with the minimum of 9 test-days per bull's daughter. As indicated in Tables 2 and 3, *DYD* are less variable than *YD*. *EBVs* and *DYDs* means of 305 days and

**Table 2.** Descriptive statistics for *YD*.

Yield Deviations ( <i>YD</i> )		Estimated Breeding Values ( <i>EBV</i> )		Correlation of <i>YD</i> with <i>EBV</i>
Mean	<i>STD</i>	Mean	<i>STD</i>	
-372.25	140.92	-0.31	34.83	0.50



**Figure 1.** (A) Yield deviations, and (B) *EBVs* for SCS at different stages of lactations for a cow with a positive *EBV* (cow A), a cow with *EBV* very close to zero (cow B) and a cow with a negative *EBV* (cow C).

**Table 3.** Descriptive statistics for *DYD* and correlation of *DYD* with *EBV* by number of daughters.

Number of daughters	Number of bulls	<i>DYD</i>		<i>EBV</i>		Correlation of <i>DYD</i> with <i>EBV</i>
		Mean	<i>STD</i>	Mean	<i>STD</i>	
10 - 19	172	-367.84	41.55	-2.25	38.65	0.87
20 - 49	166	-362.76	36.44	5.76	47.50	0.89
50 - 99	59	-369.16	35.82	-1.88	57.55	0.95
≥ 100	6	-403.84	46.49	-49.26	58.53	0.97
Total	403	-366.48	38.98	0.40	46.71	0.88

**Table 4.** Correlation of *DYD* with *EBV* by average number of test day records of daughters.

Average <i>TD</i> records of daughters	Number of bulls	Average no of daughters	<i>DYD</i>		<i>EBV</i>		Correlation of <i>DYD</i> with <i>EBV</i>
			Mean	<i>STD</i>	Mean	<i>STD</i>	
8 - 8.5	85	23	-357.68	38.33	3.96	46.79	0.89
8.51 - 9	283	33	-368.44	37.80	0.39	46.33	0.87
> 9	35	21	-371.93	47.24	-8.16	45.81	0.92

correlations between *EBVs* and *DYDs* stratified by bulls' birth year are shown in Figure 2. Correlation in all birth years was high and variation of *EBV* and *DYD* means were symmetric.

In the present study, validation of genetic trend was performed based on Interbull Method II for SCS records of Iranian Holsteins. Additive genetic standard deviation was 4.76 and regression coefficient of *DYD* on production years (SE) was 1.12 (0.37). Regression coefficients of *DYDs* on bull birth years, calculated for SCS, was positive and greater than  $0.01 \times SD$  ( $0.23 \times SD$ ), indicating that genetic trend was slightly overestimated. Bonaiti *et al.* (1994) indicated that when the estimate of genetic trend is unbiased, the year effect has a zero expectation, and should not significantly differ from zero. Alternatively, the year effect shows a decreasing or increasing trend when the estimate of genetic trend is underestimated or overestimated, respectively. Similar trends in daily *YDs* and *EBVs* at different stages of lactation are also reported by Mrode and Swanson (2004). *YD* provides a good indication of contributions

from yield records of the cow to her PTA (Predicted Transmitting Ability), thus, *YD* could be useful in understanding cow evaluations (Mrode and Swanson, 2004).

Calculation of *DYD* requires estimates of all fixed effects and non-genetic random effects and *EBV* of bulls' mates obtained from a genetic evaluation (Liu *et al.*, 2003). For bulls with granddaughters, *DYD* does not include all information from descendants, because information from granddaughters and sons is excluded (Van Raden and Wiggans, 1991). In *DYD* calculation, only the path of cow to sire is considered and other paths such as son to sire are ignored. Because *DYD* of one bull does not affect *DYD* of other bulls, *DYD* calculation can be done on a within-bull basis. Parental contribution to bull is irrelevant for the calculation of *DYD* of bulls (Liu *et al.*, 2003). For routine genetic evaluations, 305-day lactation *DYD* values and *DYD* lactation curves are published for bulls satisfying the requirement for official *DYD* mentioned above, in addition to lactation *EBV* and genetic lactation curves for bulls (Liu *et al.*, 2004).

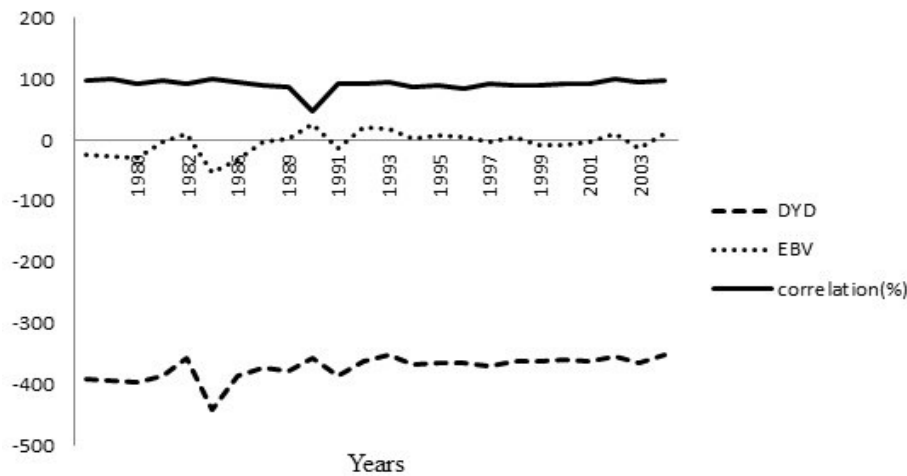


Figure 2. EBVs, DYDs and correlations between EBVs and DYDs for SCS in different bull birth years.

Increase in correlation between *DYD* and *EBV* with increase in the number of daughters per bull in this study is in agreement with results obtained by Liu *et al.* (2003, 2004), Mrode and Swanson (2004), Szyda *et al.* (2008) and Khazadeh *et al.* (2013). This is expected as

$$DYD = \frac{1}{2} \hat{a}_{bull} + \sum_m \hat{e}/m$$
, where *m* is the number of daughters. Therefore, as *m* increases,  $\sum_m \hat{e}/m$  tends towards zero and

correlation of *DYD* and *EBV* increases (Mrode and Swanson 2004). Currently, *DYD* is provided to the dairy industry for bulls with 10 or more daughters, but not for cows (Van Raden and Wiggans, 1991). Lactation *DYDs* estimated from shorter lactations are more influenced by extrapolation than lactation *DYDs* from longer lactations. To minimize the impact of extrapolation, at least 10 daughters are required to pass 120 *DIM* in lactation in order to make *DYD* of this lactation official for a bull (Liu *et al.*, 2003).

Liu *et al.* (2003), Szyda *et al.* (2008) and Khazadeh *et al.* (2013) reported that correlations between bulls' *EBVs* and *DYDs* increased by increasing average number of test-days per daughter. Szyda *et al.* (2008) suggested that a large number of test-days

(minimum 6) is required to obtain a good projection of *DYDs* over a 305-day lactation, while with increasing number of daughters per bull (over 40), the residual variance component of the total *DYD* variance strongly decreases. Since *DYDs* are considered as an additional measure of animal's genetic merit, their correlation with *EBVs* is of primary importance (Szyda *et al.*, 2008).

The large variation of *YD* is partially caused by short lactations of cows (Szyda *et al.*, 2008). On the other hand, *YD* is at cow level (based on one or few records) while *DYD* is at bull level (based on records from a large number of daughters). Therefore, variation between *YD* is larger than *DYD*. Expected *DYD* values should only depend on the bull and are theoretically independent of any environmental effect, particularly the birth year of the daughters. This property of the residuals may be used to validate the estimation of genetic trend, which is in this case the combined sire trends (Theron *et al.*, 2002). A possible reason for the observed overestimation of a trend could arise from the fact that modeling of time-related effects was not very accurate in the corresponding genetic evaluation model, e.g. age of calving or year of calving (Szyda *et al.*, 2008).

## CONCLUSIONS

*YDs* and *DYDs* were calculated for Iranian Holstein SCS records and, then, genetic trends were validated using calculated *DYDs*. Trend in daily *YDs* of cows at different stages of lactation corresponded closely with their daily *EBVs*. Calculated *DYDs* were highly correlated with *EBVs*; therefore, it can be indicated that, in addition to *EBV*, the *DYD* can be an appropriate measure for dairy cattle breeding programs. Results of the validation of genetic trend with the Interbull Method II indicated that estimate of genetic trend for SCS was slightly overestimated.

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## REFERENCES

1. Akaike, H. 1973. Information Theory and an Extension of the Maximum Likelihood Principle. *Proceedings of the 2<sup>nd</sup> International Symposium on Information Theory*, 2-8 September, Budapest, Hungary, PP. 267-281.
2. Boichard, D., Bonaiti, B., Barbat, A. and Mattalia, S. 1995 Three Methods to Validate the Estimation of Genetic Trend for Dairy Cattle. *J. Dairy Sci.*, **78**: 431-437.
3. Bonaiti, B., Boichard, D., Barbat, A. and Mattalia, S. 1994. Three Methods to Validate the Estimation of Genetic Trend in Dairy Cattle. *Interbull Meeting*, August 5-6, Ottawa, Ontario, Canada.
4. Freyer, G., Stricker, C. and Kuhn, C. 2002. Comparison of Estimated Breeding Values and Daughter Yield Deviations Used in Segregation and Linkage Analyses. *Czech J. Anim. Sci.* **47**: 247-252.
5. Haile Mariam, M., Goddard, M. E. and Bowman, P. J. 2001. Estimates of Genetic Parameters for Daily Somatic Cell Count of Australian Dairy Cattle. *J. Dairy Sci.* **84**: 1255-1264.
6. Kadarmideen, H. N. and Pryce J. E. 2001. Genetic and Economic Relationships between Somatic Cell Count and Clinical Mastitis and Their Use in Selection for Mastitis Resistance in Dairy Cattle. *J. Anim. Sci.* **73**: 229-240.
7. Khanzadeh, H., Ghavi Hossein-Zadeh, N., Naserani, M. and Mohammad Nazari, B. 2013. Calculating Daughter Yield Deviations for Production Traits in Holstein Cattle Using Repeatability Animal and Random Regression Test Day Models. *Livest. Sci.* **157**: 408-413.
8. Koivula, M., Negussie, E. and Mäntysaari, E. A. 2004. Genetic Parameters for Test-day Somatic Cell Count at Different Lactation Stages of Finnish Dairy Cattle. *Livest. Prod. Sci.* **90**: 145-157.
9. Liu, Z., Reinhardt, F., Bunger, A. and Reents, R. 2004. Derivation and Calculation of Approximate Reliabilities and Daughter Yield Deviations of a Random Regression Test-Day Model for Genetic Evaluation of Dairy Cattle. *J. Dairy Sci.* **87**: 1896-1907.
10. Liu, Z., Bunger, A. and Reents, R. 2003. Calculation and Use of Daughter Yield Deviations and Associated Reliabilities of Bulls under Multiple Trait Models. *Interbull Bull.*, **31**: 16-25.
11. Meyer K. 2006. *WOMBAT: A Program for Mixed Model Analyses by Restricted Maximum Likelihood*. User Notes. Animal Genetics and Breeding Unit, University of New England, Armidale, Australia.
12. Mrode, R. A. and Swanson, G. J. T. 1996. Genetic and Statistical Properties of Somatic Cell Count and Its Suitability as an Indirect Means of Reducing the Incidence of Mastitis in Dairy Cattle. *Anim. Breed. Abstr.*, **64**: 847-857.
13. Mrode, R. A. and Swanson, G. J. T. 2002. The Calculation of Cow and Daughter Yield Deviations and Partitioning of Genetic Evaluations when Using a Random Regression Model. *Proceedings of the 7<sup>th</sup> World Congress on Genetics Applied to Livestock Production*, 19-23 August, Montpellier, France, PP. 51-54.
14. Mrode, R. A. and Swanson, G. J. T. 2004. Calculating Cow and Daughter Yield Deviations and Partitioning of Genetic Evaluations under a Random Regression Model. *Livest. Prod. Sci.*, **86**: 253-260.
15. Ødegard, J., Jensen, J., Klemetsdal, G., Madsen, P. and Heringstad, B. 2003.



- Genetic Analysis of Somatic Cell Score in Norwegian Cattle Using Random Regression Test-day Models. *J. Dairy Sci.*, **86**: 4103-4114.
16. Reents, R., Dekkers, J. C. M. and Schaeffer, L. R. 1994. Genetic Parameters of Test Day Somatic Cell Counts and Production Traits. *Proceedings of the 5<sup>th</sup> World Congress on Genetics Applied to Livestock Production*, Guelph, Ontario, Canada, **17**: 120-123.
17. SAS. 2002. *SAS Users Guide Ver. 9.1*. Statistics SAS Institute Inc., Cary, NC.
18. Sheikloo, M., Shodja, J., Pirani, N., Alijani, S. and Sayadnejad, M. B. 2009. Genetics Evaluation and Calculating Daughter Yield Deviation of Bulls in Iranian Holstein Cattle for Milk and Fat Yields. *Asian-Austr. J. Anim. Sci.*, **5**: 611-617.
19. Szyda, J., Ptak, E., Komisarek, J. and Zarnecki, A. 2008. Practical Application of Daughter Yield Deviations in Dairy Cattle Breeding. *J. Appl. Genet.*, **49**: 183-191.
20. Theron, H. E., Kanfer, F. H. J. and Rautenbach, L. 2002. The Effect of Phantom Parent Groups on Genetic Trend Estimation. *S. Afr. J. Anim. Sci.*, **322**: 130-135.
21. Van Raden, P. M. and Wiggans, G. R. 1991. Derivation, Calculation, and Use of National Animal Model Information. *J. Dairy Sci.*, **74**: 2737-2746.
22. Van Steenberghe, E. J., Van der Lindde, C., De Roos, A. P. W., Harbers, A. G. F. and De Jong, G. 2005. Genetic Trend Validation in the PROTEJE Data and the Influence of Genetic Correlations on MACE EBVs. *Interbull Bull.*, **33**: 16-20.

## برآورد انحراف تولید دختران و اعتبارسنجی روند ژنتیکی برای امتیاز سلول های سوماتیک گاوهای هلشتاین با استفاده از مدل روزآزمون تابعیت تصادفی

ح. خانزاده، و ن. قوی حسین زاده

### چکیده

هدف از این مطالعه برآورد انحراف تولید دختران (DYDs) گاوهای نر و انحراف تولید (YDs) گاوهای ماده با استفاده از مدل تابعیت تصادفی و اعتبارسنجی روند ژنتیکی با استفاده از برآوردهای DYDs و روش دوم اینتربول برای رکوردهای روزآزمون امتیاز سلول های سوماتیک در شکم زایش اول گاوهای هلشتاین ایران بود. مجموعه داده ها مشتمل بر ۱۰۸۹۹۵ رکورد روزآزمون بود که به وسیله مرکز اصلاح نژاد دام ایران در طی سال های ۲۰۰۱ تا ۲۰۱۰ جمع آوری شده بود. نتایج این مطالعه نشان داد که YDs گاوهای ماده در مراحل مختلف شیردهی با برآورد ارزش های اصلاحی آنها (EBVs) انطباق نزدیکی دارد. چون YDs و DYDs به عنوان یک معیار اضافی از شایستگی ژنتیکی حیوان محسوب می شوند، همبستگی آنها با EBVs بسیار حائز اهمیت است. همبستگی بین DYDs و EBVs گاوهای نر برای SCS برابر با ۰/۸۸ بود. برآوردهای بالای همبستگی بین DYDs و EBVs نشان داد که DYD، علاوه بر EBV، می تواند یک معیار مناسب برای برنامه های اصلاح نژاد گاو شیری باشد. این همبستگی با افزایش تعداد دختران گاوهای نر و متوسط تعداد روزآزمون های دختران افزایش یافت. برآوردهای DYDs برای هر سال جهت تعیین اعتبار روند ژنتیکی بدست آمده از مدل ارزیابی ژنتیکی استفاده شد. نتایج نشان داد که روند ژنتیکی SCS در شکم زایش اول گاوهای هلشتاین ایران تا حد کمی بیشتر از حد معمول برآورد شدند.