Population Fluctuation and Spatial Distribution Pattern of the Nut Scale, *Eulecanium tiliae* (L.) (Hem.: Coccidae) on Cherries of the West of Iran

Z. Zareii Ahmad-Abadi¹, M. Darbemamieh¹, and H.A. Vahedi^{1*}

ABSTRACT

The nut scale, Eulecanium tiliae (L.) is a potential pest, infesting deciduous trees in Iran, where it is common in urban horticultural areas and usually present in high population on some stone fruit trees, such as Prunus spp., including Prunus cerasus. It has higher occurrence and population in sustainable orchards. This pest has one generation per year and second nymphs overwinter on one or two-years old branches. High populations of this pest can cause death of branches, twigs and degeneration of leaves chloroplasts. In this research, the spatial distribution pattern and population fluctuation for all stages of E. tiliae were investigated in a cherry orchard of Kermanshah region during two studied seasons from 2016 to 2018. Four different methods were used to calculate distribution pattern including index of dispersion, Morisita's index and regression methods (Taylor's and Iwao's). Sample size was determined for the first time, according to a primary sampling date and corrected for other sampling dates according to the data of the previous sampling date. Data analysis was performed using Minitab16 and Office Excel 2016 statistical softwares. Most used methods revealed aggregated distribution pattern of the pest in both years. The results obtained from the Morisita's index during activity showed the random distribution of the second nymphs (in 16 April 2016, Z= -1.218, $I_0 = 0.440$) (in 16 April 2017, Z= -0.179, $I_0 = 0.942$) (-1.96< Z< 1.96). These changes showed that the spatial distribution could change during the season. Knowledge of the pest spatial distribution pattern can be used to arrange sampling program in pest management.

Keywords: Iwao's Patchiness Regression, Morisita's index, Sampling program, Seasonal activity, Taylor Power Law.

INTRODUCTION

The nut scale, *Eulecanium tiliae* (Linnaeus, 1758) is very common on all fruit trees, but usually in low population. Its frequency was higher in unprotected orchards than in protected ones, amounting to 1.18 and 0.13%, respectively (Lagowska, 1984). High populations can cause death of branches, twigs, and degeneration of chloroplasts on leaves. Temperature and humidity are the main abiotic factors regulating the distribution and abundance of

soft scales (Kosztarab, 1996). Similar to other insects, developmental rate of soft scales increases in case of temperature rising until an optimal temperature is reached, after which the developmental rate declines (Abd-Rabou *et al.*, 2009).

Cosmopolitan soft scale insects may develop more than one generation in a warmer country or climatic zone within a country. In general, fecundity of soft scale species can vary enormously depending on temperature, scale abundance, body size of females and the condition of the host plant

¹ Department of Plant Protection, Campus of Agriculture and Natural Resources, Razi University, Kermanshah, Islamic Republic of ran.

^{*}Corresponding author; e-mail: Vnassah@yahoo.com

(Marotta, 1997). The mean fecundity of females in Ceroplastes sinensis Del Guercio was 3260±770 eggs per female (Stathas et al., 2003). Fecundity apiece was to 6,355 eggs for C. destructor Newstead (Wakgari and Giliomee, 2000), and 382-395 crawlers for Phalacrococcus howertoni Hodges and Hodgson (Amarasekare and Mannion. 2011). Fecundity also varies among individuals. Per capita fecundity for coccid scale, Coccus hesperidum L. counted from 70 to 1,000 eggs (Tereznikowa, 1981). Such variations might be related to the different host-plants and the different climatic conditions at the sampling sites. The growth and phenology of the scales may also be influenced by the nutrient contents of host plants (Beattie, et al., 1990).

Although diagnosing all interactions among individuals of a pest population is difficult, it is possible to identify the pattern of pest distribution as well as the changes in biological traits of a species after changes in population density. The distribution pattern is an intrinsic feature of the organisms and results from the interaction between the species and their habitat or environment and, may reflect behavioral characteristics of the species in an ecosystem (Taylor, 1984; Kuno, 1991; Moradi-Vajargah et al., 2011). The spatial distribution pattern of arthropods can provide informative description of a population and regulates the sampling program, methods of data analysis and decision-making (Iwao, 1968; Southwood and Henderson, 2000; Khaing et al., 2002). Therefore, the exact monitoring is needed in IPM tactics by performing proper sampling programs and estimating population density, which are prerequisites of effective pest control (Pedigo and Buntin, 1994; Castle and Naranjo, 2009). In most cases, spatial distribution, as one of the most important characteristic properties of insect populations in ecological communities, allows us to define them and is a typical quality in insect populations (Debouzie and Thioulouse, 1986). Knowledge of the spatial distribution provides useful information not only in theoretical population biology but also in field monitoring programs and ecology of insects (Binns *et al.*, 2000; Trumble *et al.*, 1987).

Sampling program as decision-making tool in pest management strategies plays an essential role, and spatial distribution is an integral part in sampling design (Boeve and Weiss, 1988). A change in the aggregation of a species needs an alternation for the sampling plan to have an accurate population count (Slone and Croft, 1998). Spatial distribution of each insect signifies its innate features formed by behavioral and environmental factors (Pedigo and Zeiss, 1996). Understanding the spatial distribution patterns of the insect population provides some information about behavioral and environmental factors of the population (Southwood and Henderson, 2000).

Common methods to describe spatial distribution of insect populations have been summarized by Southwood and Henderson (2000). There are different methods to calculate spatial distribution pattern including the index of dispersion (e.g., Variance-to-mean ratio. Llovd's mean crowding, Morisita's coefficient, Cole's index of dispersion, David and Moore's index, Green's index, Coefficient of 'K'), Iwao's Patchiness Regression, Taylor's Power Law, (Sedaratian et al., 2010; Darbemamieh et al., 2011). The use of dispersion indices seems to be convenient decision-making methods for management programs because of their simple procedure of calculation (Darbemamieh et al., 2011). Among the various methods, Iwao's Patchiness Regression and Taylor's Power Law were more accurate than the others to estimate the distribution pattern of the insects (Khodayari et al., 2010; Rahmani et al., 2010). In comparative studies of models, Taylor's Power Law model was mostly found to fit the spatial dispersion better than Iwao's model (Celini and Vaillant, 2004; Kapatos et al., 1996), although both methods were used to model the relationship between the mean and the variance of different arthropods. The findings of spatial distribution (i.e., regular, random or

aggregated) can determine which sampling program must be carried out, especially in sequential sampling (Feng et al., 1993). information Having about spatial distribution, density and changes in population of E. tiliae during two years, identification of factors affecting population fluctuations and determination of their effects will be helpful in management of this pest. Rare information in the literature concerning the seasonal activity of this pest led us to conduct this study in Kermanshah Province (Sahneh Region) of Iran where there is no report of a similar research.

MATERIALS AND METHODS

Sampling Procedure

Field studies were conducted in an infested cherry orchard of Sahneh (Kermanshah Province, Iran; 34° 29' 09" N, 47° 41' 29" E, Altitude: 1,376.5 m) covering an area of 300 trees ha⁻¹. The trees were about 10-years old. In this study, randomly selected 20 cm of cherry branch ends were used as a sampling unit to avoid biased estimate of population mean. Samplings were undertaken every 10 days from spring to autumn (April-October), and monthly from March 2016 till March 2018. In the laboratory, the number of motile stages of E. tiliae was counted on the same day using a stereomicroscope. Parasitized scales were put separately in a bottle with net lid and kept in 25°C incubator to rear parasitoid wasps.

Spatial Distribution

The spatial distribution of *E. tiliae* was determined by the following five methods: index of dispersion, Morisita's coefficient of dispersion, Lloyd's mean crowding, and regression techniques of Taylor's Power Law, and Iwao's Patchiness Regression.

Index of dispersion

Dispersion of a population can be determined by calculating the variance to mean ratio as follows: $S^2/m> 1$ is classified as aggregated distribution, while $S^2/m= 1$ or < 1 are classified as random or regular distribution, respectively. Departure from a random distribution can be tested by calculating the Index of Dispersion (I_D), in which *n* is the number of samples as in Equation (1):

$$I_{D} = \chi^{2} = \frac{s^{2}(n-1)}{m}$$
(1)

 I_D is approximately distributed as x^2 with n-1 degrees of freedom. Values of I_D , which fall outside a confidence interval bounded with n-1 degrees of freedom and selected probability levels of 0.95 and 0.05, for instance, would indicate a significant departure from a random distribution, and n is the number of samples. In order to test the goodness of fit, Z coefficient should be calculated according to Equation (2) shown below:

$$Z = \sqrt{2\chi^2} - \sqrt{2(\nu - 1)}$$
 (2)

Where, v is degrees of freedom (n-1). If $1.96 \ge Z \ge -1.96$, the spatial distribution will be random, but in case of z > 1.96 and z < -1.96 this parameter will be aggregative and uniform, respectively (Patil and Stiteler, 1974).

Regression Techniques

According to Taylor's Power Law, population variance (S^2) is proportional to a fractional power of the arithmetic mean (m) as in Equation (3) (Taylor, 1961):

$$\log s^2 = \log a + b \log m \tag{3}$$

Where, *a* is sample size-related scaling factor and slope b is index of aggregation which in turn recalls uniform (b< 1), random (b= 1) and aggregated (b> 1) dispersion of population (Taylor, 1961).

Iwao's Patchiness Regression method quantifies the relationship between mean

(8)

crowding index (m^*) and mean (m) using Equation (4) (Iwao, 1968):

(4)

$$m^* = \alpha + \beta m$$

Where, α indicates the tendency to crowding (positive) or repulsion (negative) and β reflects the distribution of population on space and is interpreted in the same manner as b of Taylor's Power Law (Iwao, 1968). Calculated values are compared with tabulated t-values with n-2 degrees of freedom.

Morisita's Coefficient of Dispersion $\begin{pmatrix} I_{\delta} \end{pmatrix}$ Morisita (1962) proposed a hypothesis for

testing the distribution coefficient of I_{δ} , which is calculated using Equation (5) (Pedigo and Buntin, 1994):

$$I_{\delta} = \frac{n \sum x_i (x_i - 1)}{N(N - 1)} \tag{5}$$

Where, n= The number of sample unites, x_i = The number of individuals in each sample unit and N= Total number of individuals in n samples.

Equation (6) can be used to determine significant differentiation of population from the random distribution:

$$Z = \frac{(I_{\delta} - 1)}{\left(\frac{2}{nm^2}\right)^{\frac{1}{2}}}$$
(6)

Random spatial distribution will be in case of $1.96 \ge Z \ge -1.96$, but Z < -1.96, Z > 1.96indicates regular and aggregated distribution, respectively (Pedigo and Buntin, 1994).

Lloyd's Mean Crowding (m^*)

Theoretically, mean crowding (m^*) is the mean number of other individuals per individual in the same quadrate as given by Equation (7) (Lloyd, 1967):

$$m^* = m + \left(\frac{s^2}{m} - 1\right) \tag{7}$$

Similar to variance to mean ratio, index of patchiness expressed as the ratio of mean crowding to the mean. As with the varianceto-mean ratio, the index of patchiness expressed as the ratio of mean crowding to the mean is dependent upon quadrate size, $x^*/m= 1$: Random, < 1: Regular and > 1: Aggregated (Lloyd, 1967).

Optimum Number of Sampling Units

Finding out the generic coefficients eliminates experimental needs for large sample size (Ifoulis and Savopoulou-Soultani, 2006). The optimum sample size, i.e., the smallest number of sample units with precision of estimates calculated by using coefficients a and b within Taylor's Power Law in Equation (8):

$$_{pt} = a \left(\frac{t_{\alpha/2}}{D}\right)^2 \left(\mu^{b-2}\right)$$

Where, N_{opt} = Sample size, $t_{a/2}$ = t-student of table, μ = Mean density, *a* and *b*= Taylor's coefficients and *D*= The range of accuracy. The optimum sample size is derived from Equation (9):

$$N_{opt} = \left(\frac{t_{\alpha/2}}{D}\right)^2 \left(\frac{1}{\mu} + \frac{1}{k}\right)$$
(9)

This estimation is done by Iwao's coefficients (a and b) in Equation (10) (Wilson, 1985):

$$_{opt} = \left(\frac{t_{\alpha/2}}{D}\right)^2 \left(\frac{\alpha+1}{\mu} + (\beta-1)\right)$$
(10)

Correlation coefficient values are tested for departure from zero by using the two tailed t test, at n-2 degrees of freedom.

The *D* represents the Desired fix proportion of the mean. In case of D=0.20, sample mean may be 20% higher or lower than the actual mean 95% of the time.

RESULTS AND DISCUSSION

The nut scale, *E. tiliae*, has one generation per year and overwinters as second nymph instar. The number of generation varies in different climates based on temperature and humidity and some soft scales, such as *Ceroplastes rubens* (Maskell) are univoltine in Japan and China (Itioka and Inoue, 1991; Xia *et al.*, 2005) or bivoltine in Australia (Loch and Zalucki, 1997) that may produce

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second generation, in warmer geographical regions.

Our observations showed that high population densities occasionally infested Prunus spp. in western parts of Iran. Examination of infested cherry twigs with stereomicroscope (samples) and observation of the eggs within the brood chamber in the laboratory showed that oviposition started in mid-April, peaked its population density in mid-May and ended in early June. First nymph instars emergence started at the second decade of May, maximized at late June and declined in mid-September. For the second nymph instar, early July and late August was recorded for the beginning of emergence and peak of its population density, respectively. Adults appeared in the first days of April, peaked in early May and disappeared in mid-June (Figure 1).

Both regression methods fit the data well for all the examined stages (Table 1). Evaluation of Taylor's index for all stages of nut scale *E. tillae* showed that slopes of regression line in Taylor's model and t-test were significantly larger than 1 in both 2016 and 2017, that means, spatial distributions were aggregated in both years. Taylor's regression of log S^2 on log *x* provided a good fit to the data from different *E. tiliae* stages and the values of r^2 ranged from 0.852 to 0.988 (Table 1). For all stages, the regression between log S^2 and log m was significant in Taylor's model (P< 0.05) for both years (Table 1), and can be the effect of host plant or species characters.

Data demonstrated good fitness to both Taylor's ($r^2 = 0.965$), and Iwao's model ($r^2 =$ (0.924) for 1^{st} instar, also better fitness with Taylor's model for 2^{nd} instar ($r^2 = 0.989$); and female ($r^2 = 0.972$). Also Iwao's model had good fitness for 2^{nd} instar ($r^2 = 0.973$); and female ($r^2 = 0.948$) during 2016. During 2017, Taylor's ($r^2 = 0.953$) and Iwao's model $(r^2 = 0.962)$ for 1st instar had good fitness and even better fitness with Taylor's model for 2^{nd} instar ($r^2 = 0.988$) as well as female ($r^2 =$ 0.954). It was also observed for Iwao's model for 2^{nd} instar ($r^2 = 0.948$) and female $(r^2 = 0.947)$ (Table 1). Considering feeding status, ovipositional manner, different stages behaviors and presence of stages on the host leaves, the respective model can clearly describe the observed data. This aggregation was caused by two main reasons: (1)



Figure 1. Population fluctuation curve of the different biological stages of *Eulecanium tiliae* at sampling dates from 2016 to 2018.

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Table	Iwao

		Taylor	's Power Law			
		V	В	t_b	r ²	P-value
	Egg	1.028±0.601 (-0.516-2.574)	1.594 ± 0.245 ($1.012-2.225$)	6.50	0.894	0.001^{**}
March/2016 To	1 st Instar	0.507±0.162 (0.157-0.857)	1.688 ± 0.089 ($1.496-1.881$)	18.94	0.965	0.000^{**}
March/2017	2 nd Instar	0.331 ± 0.049 (0.227-0.436)	1.877 ± 0.043 ($1.786-1.968$)	43.13	0.989	0.000**
	Female	$0.488 \pm 0.130 \ (0.180 - 0.798)$	$1.692\pm0.109(1.433-1.951)$	15.46	0.972	0.000^{**}
	Egg	1.173±0.603 (-0.303-2.650)	1.556 ± 0.264 ($1.920-2.204$)	5.89	0.852	0.001**
March/2017 To	1 st Instar	$0.878\pm0.149(0.557-1.120)$	1.423 ± 0.084 ($1.243-1.604$)	16.91	0.953	0.000**
March/2018	2 nd Instar	0.183 ± 0.055 (0.068-0.299)	1.940 ± 0.048 ($1.839-2.043$)	39.97	0.988	0.000^{**}
	Female	0.045±0.209 (-0.541-0.451)	$2.036\pm0.168(1.639-2.434)$	12.12	0.954	0.000^{**}
	Iwao's Patchiness Reg	gression				
	Egg	$78.401\pm0.588(-202.087-358.821)$	$1.808\pm0.085(1.001-2.637)$	5.61	0.863	0.002**
March/2016 To	1 st Instar	$23.880 \pm 7.919(0.098 - 47.659)$	1.492 ± 0.049 ($1.236-1.748$)	12.60	0.924	0.000^{**}
March/2017	2 nd Instar	2.343±1.099(-2.842-7.528)	2.270±0.016 (2.093-2.447)	26.70	0.973	0.000**
	Female	5.504 ± 2.073 (- $3.072-14.081$)	1.862 ± 0.045 ($1.472-2.252$)	11.30	0.948	0.000**
	Egg	12±48.87 (-318.843-342.755)	$2.439\pm0.060(1.409-3.47)$	5.79	0.848	0.001**
March/2017 To	1 st Instar	$18.668\pm5.178(-5.367-8.297)$	1.346 ± 0.037 ($1.929-2.403$)	18.89	0.962	0.000^{**}
March/2018	2 nd Instar	1.465±1.477 (-5.367-8.297)	2.166 ± 0.022 (1.929-2.403)	19.06	0.948	0.000**
	Female	4.593±2.324 (-4.587-13.774)	1.735 ± 0.048 ($1.368-2.103$)	11.18	0.947	0.000**
* Significant at $\alpha =$	0.05, ** Significant at o	r= 0.01.				

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Disharmony between habitats and environment and (2) Behaviors and factors not dependent on environmental conditions (Table 1). Compared to Iwao procedure, Taylor's Power Law mainly results in reduction of sample size measurements (Darbemamieh et al. 2011; Ifoulis and Savopulou-Soultani, 2006); and fits better in showing the distribution pattern (Afshari et al., 2009). In comparison with Iwao's method, the obtained r^2 from Taylor's model showed higher correlations with data. Therefore, it is better to use Talyor's index for the data on E. tillae, since it seems to be more appropriate for determining the type of the pest population distributions (Table 1). Taylor's Power Law as well as Iwao's Patchiness Regression have been extensively used in many arthropods for evaluation of spatial distribution, normalizing the data for statistical analysis, and sampling procedures (Davis, 1994; Deligeorgidis et al., 2002). Taylor's Power Law index must be estimated in priority and then used in practical studies. This is practically done by fitting the model to data set of estimated means and variances (Ifoulis and Savopoulou-Soultani, 2006).

Based on Lloyd's distribution index to the mean values, variance to mean ratios and values of Z for E. tillae were higher than 1.96 for both years, indicating aggregated distribution (Table 2). (S^2/m) , Index of Dispersion (I_D) , 1/k and Z test is presented in Table 3. There was a significant relationship between the mean crowding and the density of E. tiliae (P< 0.05) based on Iwao's model of March 2016 to March 2017 (Table 4).

Based on the index of dispersion, the spatial distribution in all stages was aggregated during the two study years. The Index of Dispersion (I_D) and the m^*/m values for all populations were significantly greater than 1 (Table 2), which means that this species exhibited aggregated behavior in habitat, and presence of a scale insect individual at one place may increase the probability of individual another existance nearby. Comparing 1/k values among the different stages of this species showed that the aggregation in egg stage was bigger than

able 2. Estim om March/20	ated parameters by Lloy- 16 to March/2018.	d mean crowdi	ng, index of d	lispersion, Llo	oyd mean crov	vding to mean	and common	k for Euleco	mium tiliae
	Species	ш	S ²	m*	S^2/m	I_D	Ζ	m*/m	1/k
	Egg	68.033	45763.82	739.698	672.664	1419995	1620.266	10.872	9.872
larch/2016	1 st Instar	37.884	5507.162	182.251	145.367	306870.3	718.454	4.810	3.810
0 Iomoh/2017	2 nd Instar	13.769	1179.838	98.452	85.682	180875.7	536.496	7.149	6.149
	Female	5.199	258.155	53.846	49.647	104804.9	392.870	10.355	9.355

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0.044

066.601

13268.9

520.173

3.349 9.355

> 886.408 667.459

905890

939.972 159.265

57348.26 1536.596 046.159

65.505 36.719

0.355 4.349

March/2017

3.337 5.876

1.3375.876

268959.4 71781.8

123.545 875.466

78.907 52.029

91.165 56.671

293.545

13.258 5.641

2nd Instar

March/2018

Female

Instar

Egg

March/2017

	Egg		1 st Inst	tar	2 nd Ins	tar	F	emale
Date	I_{δ}	z	I_{δ}	z	I_{δ}	z	I_{δ}	z
29/March/16	-	-	-	-	1.023	2.798	-	-
10/April/16	-	-	-	-	1.254	2.593	-	-
16/April/16	-	-	-	-	0.440	-1.218	1.77	17.886
25/April/16	2.045	1200.930	-	-	-	-	2.296	154.970
29/April/16	2.303	1909.163	-	-	-	-	1.953	136.946
05/May/16	2.427	3363.032	-	-	-	-	1.908	193.175
08/May/16	1.918	2121.853	2.548	37.479	-	-	2.360	218.865
12/May/16	1.844	2501.975	1.994	86.612	-	-	1.814	142.700
21/May/16	1.558	982.643	1.669	102.326	-	-	2.856	115.347
06/June/16	2.953	987.241	1.872	491.011	-	-	2.799	92.387
15/June/16	-	-	1.701	447.975	-	-	3.864	21.440
26/June/16	-	-	1.492	379.501	-	-	-	-
06/July/16	-	-	1.667	521.085	-	-	-	-
15/July/16	-	-	1.591	441.120	2.536	99.772	-	-
27/July/16	-	-	1.571	419.128	2.458	236.610	-	-
06/August/16	-	-	1.439	268.019	2.040	230.626	-	-
12/August/16	-	-	1.929	595.341	1.857	154.564	-	-
20/August/16	-	-	2.035	503.336	2.596	492.570	-	-
26/August/16	-	-	2.391	545.349	2.216	553.398	-	-
03/September/16	-	-	2.484	351.353	2.032	325.773	-	-
11/September/16	-	-	2.795	239.508	2.699	288.083	-	-
23/September/16	-	-	-	-	2.402	252.766	-	-
03/October/16	-	-	-	-	2.786	274.005	-	-
17/October/16	-	-	-	-	2.962	235.774	-	-
02/November/16	-	-	-	-	3.051	133.408	-	-
18/November/16	-	-	-	-	2.474	81.131	-	-
27/November/16	-	-	-	-	2.661	47.274	-	-
14/December/16	-	-	-	-	3.410	23.708	-	-
01/January/17	-	-	-	-	2.133	4.438	-	-
04/February/17	-	-	-	-	1.454	2.372	-	-
26/February/17	-	-	-	-	1.761	4.438	-	-
12/March/17	-	-	-	-	2.409	4.416	-	-

Table 3. Morisita's index and Z values for *Eulecanium tiliae* in different sampling dates from March/2016 to March/2017 (Z test for goodness of fit).

that of the other stages because higher value of the 1/k index that might be due to aggregated oviposition under scales. All of the calculated 1/k values suggest aggregated pattern of dispersion for both years in all stages (Table 2).

There were some differences in Morisita's index values of each stage, but in most sampling dates, the index was significantly more than one and Z was greater than 1.96 (Tables 3 and 4), suggesting that the spatial distribution of all stages was aggregated. The results obtained from Morisita's index when *E. tillae* was present and active were

significantly higher than one for all dates, except for the second instar nymphs. The spatial distributions were random on the dates of 16 April 2016 and 18 April 2017. This was because the second instar nymphs were molting on those dates and some of them changed their sheath and became immature females (Tables 3 and 4).

Since Morisita's coefficient estimates spatial distribution using the mean and variance of each sampling date separately, this index is more accurate than the dispersion index. Showing one distribution per date, it can be used to understand details

	Egg		1 st Insta	ar	2 nd Inst	ar	Female	
Date	I_{δ}	z	I_{δ}	z	I_{δ}	z	I_{δ}	z
26/March/17	-	-	-	-	1.26	2.576	-	_
07/April/17	-	-	-	-	1.13	1.972	-	-
18/April/17	-	-	-	-	0.942	-0.179	1.434	15.356
25/April/17	1.879	757.105	-	-	-	-	1.809	72.548
05/May/17	2.165	1647.7	-	-	-	-	1.730	172.532
09/May/17	1.766	2119.799	-	-	-	-	1.721	190.873
12/May/17	2.676	4916.712	3.390	100.072	-	-	2.199	209.469
15/May/17	3.24	5332.203	2.408	130.553	-	-	2.411	120.391
19/May/17	2.178	1992.21	1.583	78.162	-	-	2.4003	126.152
28/May/17	1.411	132.787	1.565	110.717	-	-	2.267	81.994
07/June/17	17.21	835.740	1.469	200.735	-	-	1.607	17.774
16/June/17	-	-	1.394	274.110	-	-	-	-
27/June/17	-	-	1.423	321.601	-	-	-	-
07/July/17	-	-	1.561	405.851	-	-	-	-
16/July/17	-	-	1.591	441.120	2.155	63.344	-	-
28/July/17	-	-	1.556	407.980	2.13	160.398	-	-
07/August/17	-	-	1.401	244.737	1.815	72.099	-	-
13/August/17	-	-	1.380	243.741	1.685	108.961	-	-
21/August/17	-	-	1.554	299.263	1.847	253.986	-	-
27/August/17	-	-	1.912	372.719	1.991	445.768	-	-
04/September/17	-	-	2.839	320.436	2.779	545.028	-	-
12/September/17	-	-	4.049	123.972	2.619	292.001	-	-
24/September/17	-	-	-	-	2.004	299.214	-	-
04/October/17	-	-	-	-	2.575	262.511	-	-
18/October/17	-	-	-	-	2.92	232.504	-	-
03/November/17	-	-	-	-	2.207	201.942	-	-
19/November/17	-	-	-	-	3.033	132.905	-	-
28/November/17	-	-	-	-	2.738	26.632	-	-
15/December/17	-	-	-	-	2.433	23.951	-	-
31/December/17	-	-	-	-	1.605	9.638	-	-
05/February/18	-	-	-	-	1.488	2.424	-	-
27/February/18	-	-	-	-	1.846	4.860	-	-
13/March/18	-	-	-	-	1.770	3.152	-	-

Table 4. Morisita's index and Z values for *Eulecanium tiliae* in different sampling dates from March/2017 to March/2018 (Z test for goodness of fit).

of dispersion in different sampling dates that would be useful for research strategies more than management programs. It shows that distribution pattern can change as the result of some factors during the season, and distribution pattern in most of the sampling dates can be considered as a base for management decisions. Although Taylor's indices have been widely used by many researchers, we suggest Morisita's index for research programs because of its higher determination coefficient as well as better dispersion interpretation for these species (Tables 3 and 4). Different statistical methods have various accuracies in calculating spatial distribution of an organism.

One of the major reasons for having aggregate distributions is the proximity between the place of aggregated eggs and the place of feeding, which provides larva with sufficient and proper food, suggesting that such a pattern of spatial distribution is of the reproductive type distributions (Taylor, 1961). The study on the spatial distributions of 3 species of scale insects indicated that phenology and the part of the host plant had no effects on intra-species distributions, and levels of aggregation index were fixed within a species. For those three species of studied scales, spatial distributions were aggregated (Nestel *et al.*, 1995). Coefficients a and b are related to the host plants and appropriateness of some parts of habitats in addition to being dependent on insect species (Jones, 1990; Raworth, 1986). In general, it appears that spatial distribution patterns of populations are caused by innate behavior patterns, morphological features of host plants, and

Table 5. Calculated sample size of egg, 1^{st} instar, 2^{nd} instar and female populations on cherry leaves based on *k* in negative binomial distribution and Taylor's Power Law and Iwao's Patchiness coefficients from March/2016 to March/2018.

	n _{opt}	Egg	1 st Instar	2 nd Instar	Female
March/2016	K	818.762	317.752	515.281	790.658
То	Taylor	18.590	16.376	24.068	29.487
March/2017	Iwao	470.120	904.526	95.880	214.289
March/2017	K	1339.142	337.127	596.358	924.053
То	Taylor	18.426	11.0135	15.779	4.834
March/2018	Iwao	3575.440	1075.581	9.019	168.947



Figure 2. Estimated optimum sample sizes for different stages of *Eulecanium tiliae* based on Taylor's coefficients in two levels of 15 and 25% accuracy for 2016 to 2017 sampling years

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overall effects of the major factors involved in their mortality (Nestel *et al.*, 1995).

In this survey, the separate counts of *E. tiliae* stages were used to develop generic coefficients of regression techniques based on large amount of data in each stage. The sample sizes for each stage were recalculated using Taylor's and Iwao's coefficient, i.e. *a*, *b*, α and β (Table 5). The lowest estimated sample sizes for different stages were calculated by using Taylor's Equation, for all species. For example, when the mean density of nit population was 50 insects per sample unit, optimal sample number obtained by Taylor's method with precision of 15% and 25% were 35 and 15, respectively, in 2017, but the obtained figures by Iwao's method with the same precision in the same year were 130 and 50, respectively (Figures 2 and 3). According to Figures 2 and 3, optimum sample size for any of different pest stages varies by a specified mean, after which optimum sample size becomes fixed for each stage.

In comparative studies of models, Taylor's Power Law model usually showed to fit the spatial dispersion better than Iwao's model and achieved an appropriate precision of estimates (Afshari *et al.*, 2009). Taylor's method reduces the necessary sample size



Figure 3. Estimated optimum sample sizes for different stages of *Eulecanium tiliae* based on Iwao's coefficients at two levels of 15 and 25% accuracy for 2016 to 2018 sampling years.

by almost half compared to the Iwao's method. In order to acquire greater precision, the 15% level was adopted, whereas in IPM programs 25% level is acceptable. The optimum number of samples suggested by Taylor's lines, taking into account a desired accuracy of 25%, is typically higher at low population level. The information of spatial distribution can determine sampling program, especially in sequential sampling. The use of dispersion indices seems to be a convenient decisionmaking method for management programs because of the easy calculation procedure and simple results. The result of this study can be used in places with similar climate.

REFERENCES

- Abd-Rabou, S., Ali, N. and El-Fatih, M. M. 2009. Life Table of the Hemispherical Scale, *Saissetia coffeae* (Walker) (Hemiptera: Coccidae). *Egypt. Acad. J. Biolog. Sci.*, 2: 165–170.
- Afshari, A., Soleyman-Negadian, E. and Shishebor, P. 2009. Population Density and Spatial Distribution of *Aphis gossypii* Glover (Homoptera: Aphididae) on Cotton in Gorgan, Iran. J. Agr. Sci. Tech., 11: 27-38.
- Amarasekare, K. G. and Mannion, C. M. 2011. Life History of an Exotic Soft Scale Insect *Phalacrococcus howertoni* (Hemiptera: Coccidae) Found in Florida. *Fla. Entomol.*, 94: 588–593.
- Beattie, G. A. C. and Kaldor, C. J. 1990. Comparison of High-Volume Oscillating Boom and Low-Volume Fan-Assisted Rotary Atomizer Sprayers for the Control of Chinese Wax Scale, *Ceroplastes sinensis* Del Guercio (Hemiptera: Coccidae), on Valencia orange, Citrus sinensis (L.) Osbeck. J. Appl. Entomol., 22: 49–53.
- Binns, M. R., Nyrop, J. P. and Werf, W. 2000. Sampling and Monitoring in Crop Protection: The Theoretical Basis for Developing Practical Decision Guides. CABI, UK.
- 6. Boeve, P. J. and Weiss, M. 1988. Spatial Distribution and Sampling Plans with Fixed Levels of Precision for Cereal Aphids (Homoptera: Aphididae) Infesting Spring Wheat. *Can. Entomol.*, **130**: 67-77.

- Castle, S. J. and Naranjo, S. E. 2009. Sampling Plans, Selective Insecticides and Sustainability: The Case for IPM as 'Informed Pest Management'. *Pest. Manag. Sci.*, 65: 1321-1328.
- Celini, L. and Vaillant, J. 2004. A Model of Temporal Distribution of *Aphis gossypii* Glover (Homoptera: Aphididae) on Cotton. *J. Appl. Entomol.*, **128**(2): 133-138.
- Darbemamieh, M., Fathipour, Y. and Kamali, K. 2011. Population Abundance and Seasonal Activity of *Zetzellia pourmirzai* (Acari: Stigmaeidae) and Its Preys *Cenopalpus irani* and *Bryobia rubrioculus* (Acari: Tetranychidae) in Sprayed Apple Orchards of Kermanshah, Iran. J. Agr. Sci. Tech., 13: 143-154.
- Davis, P. M. 1994. Statistics for Describing Populations. In: "Handbook of Sampling Methods for Arthropods in Agriculture", (Eds): Pedigo, L. P. and Buntin, D. G. CRC Press, Boca Raton, Florida, PP. 35-54.
- Deligeorgidis, P. N., Athanassiou, C. G. and Kavallieratos, N. G. 2002. Seasonal Abundance, Spatial Distribution and Sampling Indices for Trips Populations on Cotton: Four-Years Survey from Central Greece. J. Appl. Entomol., 126: 343-348.
- Debouzie, D. and Thioulouse, J. 1986. Statistics to Find Spatial and Temporal Structures in Populations. In: "Pest Control: Operations and System Analysis in Fruit Fly Management", (Ed.): Mangel, M. NATO ASI Series, Springer, Berlin, G11: 263-282.
- Feng, M. G. and Nowierski, R. M. 1992 a. Spatial Distribution and Sampling Plans for four Species of *Cereal Aphids* (Homoptera: Aphididae) Infesting Spring Wheat in Southwestern Idaho. *J. Econ. Entomol.*, **85**(3): 830-837.
- Feng, M. G. and Nowierski, R. M. 1992b. Variation in Spatial Pattern of the Russian Wheat Aphid (Homoptera: Aphididae) among Small Grains in the Northwestern United States. *Environ. Entomol.*, 21(5): 1029-1034.
- Feng, M. G., Nowierski R. M., Zeng, Z. and Scharen, A. L. 1993. Estimation of Population Density of the Russian Wheat Aphid (Homoptera: Aphididae) from the Proportion of Grain Tillers with the Different Tally Threshold of Aphids. J. Econ. Entomol., 86: 427-435.
- 16. Ifoulis, A. A. and Savopulou-Soultani, M. 2006. Developing Optimum Sample Size

and Multistage Sampling Plans for *Lobesia botrana* (Lepidoptera: Tortricidae) Larval Infestation and Injury in Northern Greece. J. *Econ. Entomol.*, **99**: 1890-1898.

- Itioka, T., and Inoue T. 1991. Settling-Site Selection and Survival of Two Scale Insects, *Ceroplastes rubens* and *C. ceriferus*, on Citrus Trees. *Res. Popul. Ecol.*, 33: 69–85.
- Iwao, S., 1968. A New Regression Method for Analyzing the Aggregation Pattern of Animal Populations. *Res. Popul. Ecol*, **10**: 1-20.
- Jones, V. P. 1990. Developing Sampling Plans for Spider Mites (Acari: Tetranychidae): Those Who Don't Remember the Past May Have to Repeat It. *J. Econ. Entomol.* 83: 1656–1664.
- Lagowska, B. 1984. Badania nad Nasileniem wyst (Powania miseczników (Homoptera, Coccidae) na Drzewach i Krzewach owocowych w Okolicach lublina. Roczniki Nauk Rolniczych. Seria, 14:9-17.
- Loch, A. D. and Zalucki M. P. 1997. Variation in Length, Fecundity and Survival of Pink Wax Scale, *Ceroplastes rubens* Maskell (Hemiptera: Coccidae), on Umbrella Trees. *Aust. J. Zool.*, 45: 399–407.
- Southwood, T. R. E. and Henderson, P. A. 2000. Ecological Methods. Third Edition, Blackwell Sciences, Oxford, 592 PP.
- 23. Kapatos, E. T., Stratopoulou, E. T., Sahinoglou, A., Tsitsipis, J. A. and Lycouresis, D. P. 1996. Development of an Optimum Sampling Plan for the Population of *A. gossypii* (Homoptera: Aphididae) on Cotton in Greece. *J. Appl. Entomol.*, **120**: 245-248.
- Khaing, O., Hormchan, P., Jamornmarn, S. and Wongpiyasatid, A. 2002. Spatial Dispersion and Optimum Sample Size for Cotton Bollworm, *Helicoverpa armigera* (Hubner) (Lepidoptera: Noctuidae) Larvae on Cotton. *Witthayasan Kasetsat*, 36: 235-241.
- Khodayari, S., Fathipour, Y., Kamali, K. and Naseri, B. 2010. Seasonal Activity of *Zetzellia mali* (Stigmaeidae) and Its Preys *Eotetranychus frosti* (Tetranychidae) and *Tydeus longisetosus* (Tydeidae) in Unsprayed Apple Orchards of Maragheh, Northwestern of Iran. J. Agr. Sci. Tech., 12: 549-558.
- 26. Kosztarab, M. 1996. Scale Insects of Northeastern North America: Identification,

Biology, and Distribution. Virginia Museum of Natural History, Special Publication Number 3, Martinsville, Virginia, USA, 650 PP.

- Kuno, E., 1991. Sampling and Analysis of Insect Population. Annu. Rev. Entomol., 36: 285-304.
- 28. Lloyd, M. 1967. Mean Crowding. J. Anim. Ecol., **36**:1-30.
- Marotta, S. 1997. General Life History. In "Soft Scale Insects: Their Biology, Natural Enemies and Control", (Eds.): Ben-Dov, Y. and Hodgson, C. J. Elsevier Science BV, Amsterdam, The Netherlands, 7A: 251–255.
- Moradi-Vajargah, M., Golizadeh, A., Rafiee Dastjerdi H., Zalucki M. P. and Hassanpour M. 2011. Population Density and Spatial Distribution Pattern of *Hypera postica* (Coleoptera: Curculionidae) in Ardabil. *Notulae Botanicae Horti Agrobotanici Cluj-Napoca*, **39(2)**: 42-48.
- Morisita, M. 1962. I_δ-Index: A Measure of Dispersion of Individuals. *Res. Popul. Ecol.*, 4(1): 1–7.
- Nestel, D. and Klein, M. 1995. Geostatistical Analysis of Leafhopper (Homoptera: Cicadellidae) Colonization and Spread in Deciduous Orchards. *Environ. Entomol.*, 24: 1032–1039.
- 33. Patil, G. P. and Stiteler, W. M. 1974. Concepts of Aggregation and Their Quantification: a Critical Review with Some New Result and Applications. *Res. Popul. Ecol.*, **15**: 238-254.
- 34. Pedigo, L. P. and Buntin, G. D. 1994. *Handbook of Sampling Methods for Arthropods in Agriculture*. CRC Press, Florida, 714 PP.
- 35. Pedigo, L. P. and Zeiss, M. R. 1996. Analysis in Insect Ecology and Management. Iowa State University Press, Ames, 168 PP.
- 36. Rahmani, H., Fathipour, Y. and Kamali, K. 2010. Spatial Distribution and Seasonal Activity of *Panonychus ulmi* (Acari: Tetranychidae) and Its Predator *Zetzellia mali* (Acari: Stigmaeidae) in Apple Orchards of Zanjan, Iran. J. Agr. Sci. Tech., **12**: 155-165.
- Raworth, D. A. 1986. Sampling Statistics and a Sampling Scheme for the Two Spotted Spider Mite, *Tetranychus urticae* (Acari: Tetranychidae), on Strawberries. *Can. Entomol.*, **118**(8): 807–814.

- Sedaratian, A., Fathipour, Y., Talebi, A. A. and Farahani, S. 2010. Population Density and Spatial Distribution Pattern of *Thrips tabaci* (Thysanoptera: Thripidae) on Different Soybean Varieties. J. Agr. Sci. Tech., 12: 275-288.
- 39. Slone, D. H. and Croft, B. A. 1998. Spatial Aggregation of Apple Mites (Acari: Phytoseiidae, Stigmaeidae, Tetranychidae) as Measured by a Binomial Model: Effects of Life Stage, Reproduction, Competition and Predation. *Environ. Entomol.*, 27: 918-925.
- 40. Southwood, T. R. E. 1995. Ecological Methods, with Particular Reference to the Study of Insect Population. Chapman and Hall, London.
- 41. Stathas, G., Kavallieratos, N. and Eliopoulos, P. 2003. **Biological** and Ecological Aspects of Chinese Wax Scale, Ceroplastes sinensis Del Guercio (Hemiptera: Coccidae): A Two-Year Study from Central Greece. Aust. J. Entomol., 42: 271-275.
- 42. Taylor, L. R. 1961. Aggregation, Variance to the Mean. *Nature*, **189**: 732-735.
- 43. Taylor, L. R. 1984. Assessing and Interpreting the Spatial Distributions of

Insect Populations. *Annu. Rev. Entomol.*, **29**: 321-357.

- 44. Tereznikowa, E. M. 1981. Scale Insects: Eriococcidae, Kermesidae and Coccidae. Fauna Ukraini. Academiya Nauk Ukrainskoi RSR. *Institut Zoologii. Kiev*, **20**: 1–215.
- 45. Trumble, J. T., Edelson, J. V. and Story, R. N. 1987. Conformity and incongruity of Selected Dispersion Indices in Describing the Spatial Distribution of *Trichoplusia* ni (Hubner) in Geographically Separate Cabbage Plantings. *Res. Popul. Ecol.*, 29: 155-166.
- 46. Wakgari, W. M. and Giliomee, J. H. 2000. Fecundity, Fertility and Phenology of White Wax Scale, *Ceroplastes destructor* Newstead (Hemiptera: Coccidae), on Citrus and Syzygium in South Africa. *Afr. Entomol.*, 8: 233–242.
- 47. Wilson, J. R. 1985. *Introduction to Graph Theory*. 3rd Edition. Longman Scientific and Technical, Essex, UK.
- Xia, C. Y., Zhang, W., Sun, X. Q., Li, H. P. and Dai, G. H. 2005. Observations on Biological Habits of *Ceroplastes rubens* Maskell in Shanghai. J. Shanghai Jiaotong Univ. J. Agric. Sci., 4: 022.

نوسانات جمعیت و الگوی توزیع فضایی شپشک نخودی (L.) Eulecanium tiliae (Hem.: Coccidae) روی گیلاس در غرب ایران

ز. زارعی احمد آبادی، م. درب امامیه، و ح. ع. واحدی

چکیدہ

شپشک نخودی (Hem.: Coccidae)، در اکثر مناطق ایران به عنوان یک آفت بالقوه برای انواع درختان میوه است و در مناطق عمده باغداری در ایران روی هسته داران . *Prunus* spp و از جمله *Prunus cerasus* به شمار می آید. این آفت حضور و جمعیت بیشتری در باغات سمپاشی نشده دارد. این بررسی نشان داد که این آفت یک نسل در سال دارد و زمستان گذرانی آن به صورت پوره سن دوم بر روی سرشاخههای یک یا دو ساله است. مشاهدات نشان داده که آلودگی شدید باعث زرد شدن زودرس برگ و ضعف درختان می شود. در این تحقیق ، الگوی توزیع فضایی ، نوسانات جمعیت و تغییرات فصلی برای تمام مراحل *E. tiliae* در یک باغ گیلاس در استان کرمانشاه طی دو فصل رشد در سال های 1395 و 1396 بررسی شد. برای تعیین الگوی توزیع فضایی حشرات مورد مطالعه از چهارروش شاخص پراکندگی، موریستا و روشهای رگرسیونی (تیلور و آیوائو) استفاده شد. برای تعیین تعداد نمونهی مورد نیاز، در ابتدا یک نمونهبرداری مقدماتی انجام شد و برای سایر تاریخها با توجه به داده های نمونه برداری تاریخ قبلی، تعداد نمونه مورد نیاز اصلاح شد. تجزیه و تحلیل داده ها با استفاده از نرم افزارهای آماری Minitab16 و Microsoft office و شدا شد. تجزیه و تحلیل داده ها با استفاده از نرم افزارهای آماری 2016 مالی و Minitab16 و Microsoft office شد. تجزیه و تحلیل داده ها با استفاده از نرم افزارهای آماری 2016 مالی و Microsoft office و مراد سال تجمعی نشان داده شد. نتایج به دست آمده از شاخص موریستا نشان داد که توزیع تصادفی پوره سن دوم (در 28 فروردین 1395، 1400 = *م*ا ، 2018 – *Z*)، (در 27 فروردین 1396، - *Z* دوم (در 28 فروردین 1395، 2040) میاشد که این تغییرات نشان می دهد که توزیع فضایی در طول فصل تغییر میکند. آگاهی از الگوی توزیع فضایی آفت می تواند در تعیین بهترین تصمیم گیری جهت کنترل آفت در برنامه های مدیریت تلفیقی آفات کمک کند.

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