Improvement of Reproductive and Growth Traits by Selecting Ewes Based on Reproductive Composite Traits in Baluchi Sheep

M. Bagheri¹*, M. Moradi-Shahrbabak², R. Vaez-Torshizi³, and A. Zahmatkesh¹

ABSTRACT

In this study, 1973-2003 data of 12,328 animals from Abbasabad Baluchi Sheep Breeding Station, Iran, were used to estimate genetic and phenotypic correlations between Total Weight of lambs Weaned per Ewe Joined in the first, 2nd, 3rd and 4th lambing (TWW1/EJ, TWW2/EJ, TWW3/EJ, and TWW4/EJ, respectively), and TWW of the total four lambing (TWW/EJ). Also, the study aimed to estimate correlations of TWW/EJ...TWW4/EJ and Birth Weight (BW) as well as Weaning Weight (WW). The (co)variance components and genetic correlations were estimated by DFREML procedures. Genetic correlation between TWW1/EJ and TWWt/EJ was high and positive (0.76). Genetic correlations of TWW/EJ in different lambing with BW and WW ranged from 0.16 to 0.32 and 0.95 to 0.97, respectively. Repeatability of TWW/EJ for lambing 1-4 was 0.15. Results suggested that selection for increased lifetime reproductive performance could be based on TWW1/EJ in the first lambing. Overall, results indicate that TWW/EJ, as an important but sex-limited trait, may be used in selection criterion to indirectly improve the related traits.

Keywords: Birth weight, Composite traits, Heritability, Weaning weight.

INTRODUCTION

Baluchi sheep, a fat-tailed and small sheep, encompasses about 20% of the sheep population in Iran and can be found mostly in the arid and semi-arid regions (Vaez et al., 1993). This breed, as the most numerous native breed of Iran, is well adapted to dry and hot climatic conditions as well as low quality pasture in the eastern part of this country (Yazdi et al., 1997). Selling lambs is the most important source of income in this production system. Hence, an effective way of improving production efficiency is to increase total weight of lambs weaned per ewe in a given period as a composite trait. This composite trait can be determined by litter size as well as several other traits such as maternal ability, milk production of ewes, and growth potential of lambs.

Although in harsh conditions, with restrictions on food sources, increasing the number of born lambs may increase the total weight of lambs weaned per ewe, the quality of these lambs is not acceptable because of inappropriate nutrition. Total weight of weaned lambs is affected by many factors including numbers of born lambs, maternal capacity, ewe's milk production, and lamb's growth potential (Snyman et al., 1998a). Therefore, the genes affecting these traits may have

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effects on total weight of lamb weaned and lamb number or total weight of lamb weaned per ewe, as a criteria target in sheep breeding. This could partly be obtained by increasing the number of lambs weaned and weight of lambs weaned per ewe during a specific year (Duguma et al., 2002). Since, improvement in ewe productivity can promote meat productions, lamb production would be the major source of profitability in sheep production system (Ekiz et al., 2005). Litter size at weaning includes the number of lambs survived at weaning, but not the weight. It would be kept in mind that selection for only litter size at birth would not be effective for increasing lamb production, so it could not include the survival rate and the individual lamb’s weight at weaning (Rosati et al., 2002). Therefore, the litter weight of lambs weaned per ewe can be considered as the most important factor for ewe’s reproduction and the economic efficiency of lamb production (Mohammadi et al., 2013; Sakthivel et al., 2017).

Many research groups have reported genetic improvement in reproductive traits, but reported results are very variable, mainly due to the low heritability of these traits. Therefore, indirect selection for improving reproductive traits will be more effective than direct selection. Therefore, estimation of genetic and phenotypic correlation between traits will be necessary to construct (co)variance components and genetic and phenotypic correlations among composite reproductive and productive traits for harsh system-living breeds such as Baluchi sheep. Estimation of repeatability can be useful for prediction of producing ability of animals in their repeated records (Bourdon, 2000). The purpose of the present study was evaluation of selection based on reproductive composite traits (total weight of lamb weaned per ewe joined in the first, second, third and fourth lambing) and related production traits (birth and weaning weight).

MATERIALS AND METHODS

Data and Pedigree

Information of Baluchi sheep were collected between 1973 and 2003 (30 years) from Abbasabad Baluchi Sheep Breeding Station, Iran. The data set for all traits included 12,328 animals, 229 sires, 3671 dams, and 1428 ewes. A detailed description of management and executed selection procedures has been given by Sanaee et al. (2002). Briefly, selection was based on body weight, body conformation score, and type of birth. Mating was random for each ram with 15 to 25 ewes. In this flock, about 50% of sires were kept for 2 to 3 mating seasons and the rest were used just once, and the longevity of ewes were up to 7 lambing (about 8 years of age). Data such as sex, birth date, birth type, birth weight, sire ID and dam ID were recorded after lambing. Also, records of body weights for weaning (90 days on average), 6 and 9 months of age were registered. Flock was fed in pasture during spring and summer, but in autumn and winter, they were fed with a ration based on wheat and barley stubble, alfalfa hay, dry sugar beet pulp, and concentrate. Traits analyzed were Birth Weight (BW) and Weaning Weight (WW), and the composite traits included Total Weight of lamb Weaned per Ewe Joined in the first (TWW1/EJ), second (TWW2/EJ), third (TWW3/EJ) and fourth (TWW4/EJ) lambing and also Total Weight of lamb weaned per Ewe Joined for all four lambing (TWWt/EJ).

Statistical Analysis

At first, data were verified after deleting out-of-range data and lambs without weight records or with incomplete parentage records. For fixed effects, significance level was considered at P< 0.05. TWW1/EJ to TWW4/EJ were calculated by adding the weaning weight of lambs of each ewe in
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each lambing considering the continuity of lambing after correction for sex effect. Therefore, a limitation of data preparation for TWW1/EJ... TWW4/EJ was that the data needed to be continuous. For example, for providing TWW2/EJ data, ewes should have had TWW1/EJ and TWW2/EJ records, for TWW3/EJ data, ewes should have had TWW1/EJ, TWW2/EJ, and TWW3/EJ; and for TWW4/EJ data, ewes should have had TWW1/EJ, TWW2/EJ, TWW3/EJ, and TWW4/EJ records. This undesirably decreases the number of TWW records for the third and fourth lambing. For TWW1/EJ to TWW4/EJ, at the first step, sex differences of lambs were corrected, but the age of lambs was considered as a covariate. To avoid overestimation of the total weaning weight of twin per ewe, birth type was not corrected.

Correlation coefficients among BW, WW and TWW/EJ in different lambing were estimated using the available records in ewes. Co(variance) components and genetic and phenotypic correlations were estimated using bivariate animal models and DFREAL program (Meyer, 2000). Likelihood ratio test was used to determine the most suitable model for each trait. The following models were fitted for TWW1/EJ, TWW2/EJ, TWW3/EJ, TWW4/EJ and TWWt/EJ:

\[ y = Xb + Z_a + W_s + e \]  

(1)

\[ y = Xb + Z_d + W_t + e \]  

(2)

Where, \( y \) is the vector of observed records; \( b \) is the vector of fixed effects (year-season, age of dam as well as age of weaning weight for TWW1/EJ, TWW2/EJ, TWW3/EJ and TWW4/EJ fitted as covariates), \( a \) and \( s \) are the vectors of direct additive genetic effects and random effect of service sire effects, respectively. \( X, Z_a \text{ and } W_s \) are the corresponding incidence matrices related to \( y \), and \( e \) is the vector of residuals.

And the model for BW and WW was:

\[ y = Xb + Z_d a + Z_m m + Z_p Pe + e (Cov a, m \neq 0) \]

(3)

Where, \( y \) is the vector of observed records; \( b \) is the vector of fixed effects (year-season, sex, type of birth and age of lamb at weaning for WW fitted as a covariate), and \( a, m \text{ and } Pe \) are the vectors of direct additive genetic effects, maternal additive genetic, and maternal permanent environmental effects, respectively. \( X, Z_d, Z_m \text{ and } Z_p \) are the corresponding incidence matrices related to \( y \), and \( e \) is the vector of residuals.

The heritability (\( h^2 \)) was calculated as the ratio of animal additive genetic variance to phenotypic variance (Ramatsoma et al., 2015):

\[ h^2 = \frac{\sigma^2_a}{\sigma^2_t} \]  

(4)

Repeatability (\( r \)) for total weight of lamb weaned per ewe joined in four lambing was calculated using the following formula:

\[ r = \frac{\sigma^2_a + \sigma^2_P}{\sigma^2_P} \]  

(5)

Where, \( \sigma^2_a \), \( \sigma^2_P \) and \( \sigma^2_P \) were direct additive genetic variance, ewe permanent environmental variance, and phenotypic variance, respectively. To estimate the genetic, phenotypic, and environmental parameters and correlations between composite reproductive traits, univariate and bivariate analyses were performed, respectively. The fixed effects for both univariate and bivariate animal models were the same and included year-season, age of dam, and age of lamb at weaning for WW fitted as a covariate. For bivariate analysis, the estimation of parameters was carried out in two steps, using Powell’s method. In the first step, variance components were fixed at univariate estimates and the likelihood was maximized only with respect to the covariance, then, a second run was performed to estimate all parameters. The convergences criterion was set for the first and second steps.

**RESULTS**

The number of records, mean, and CV for each trait is presented in Table 1. The coefficient of variance and standard deviation increased with increase in number of records. Estimates of (co)variance components, heritability, variance ratio of permanent environmental effects to the phenotypic variance, and repeatability for composite reproductive traits are presented.
Table 1. Number of records, mean standard deviation, and coefficient of variance for production and composite reproductive traits in Baluchi sheep.

<table>
<thead>
<tr>
<th>Traits</th>
<th>No of records</th>
<th>Mean (kg)</th>
<th>SD</th>
<th>CV</th>
</tr>
</thead>
<tbody>
<tr>
<td>BW</td>
<td>12328</td>
<td>4.28</td>
<td>0.67</td>
<td>15.78</td>
</tr>
<tr>
<td>WW</td>
<td>11044</td>
<td>22.65</td>
<td>4.55</td>
<td>20.12</td>
</tr>
<tr>
<td>TWW1/EJ</td>
<td>2593</td>
<td>23.65</td>
<td>5.44</td>
<td>19.70</td>
</tr>
<tr>
<td>TWW2/EJ</td>
<td>1402</td>
<td>26.17</td>
<td>6.89</td>
<td>23.36</td>
</tr>
<tr>
<td>TWW3/EJ</td>
<td>808</td>
<td>27.36</td>
<td>7.75</td>
<td>26.03</td>
</tr>
<tr>
<td>TWW4/EJ</td>
<td>417</td>
<td>27.60</td>
<td>8.07</td>
<td>26.14</td>
</tr>
</tbody>
</table>

Table 2. Estimation of variance components and genetic parameters for composite reproductive traits based on univariate analysis.

<table>
<thead>
<tr>
<th>Traits</th>
<th>Model</th>
<th>$\sigma^2_a$</th>
<th>$\sigma^2_s$</th>
<th>$\sigma^2_p$</th>
<th>$h^2_a(SE)$</th>
<th>$S^2(SE)$</th>
<th>LogL</th>
</tr>
</thead>
<tbody>
<tr>
<td>TWW1/EJ</td>
<td>1</td>
<td>2.44</td>
<td>-</td>
<td>19.26</td>
<td>21.70</td>
<td>0.11(0.03)</td>
<td>-5259.4</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>2.44</td>
<td>0.000</td>
<td>19.26</td>
<td>21.70</td>
<td>0.11(0.03)</td>
<td>-5259.4</td>
</tr>
<tr>
<td>TWW2/EJ</td>
<td>1</td>
<td>1.04</td>
<td>-</td>
<td>36.33</td>
<td>37.38</td>
<td>0.03(0.03)</td>
<td>-3026.0</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>1.04</td>
<td>0.000</td>
<td>36.34</td>
<td>37.38</td>
<td>0.03(0.03)</td>
<td>-3026.0</td>
</tr>
<tr>
<td>TWW3/EJ</td>
<td>1</td>
<td>3.79</td>
<td>-</td>
<td>46.95</td>
<td>50.76</td>
<td>0.07(0.06)</td>
<td>-1953.0</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>3.51</td>
<td>2.745</td>
<td>45.14</td>
<td>51.03</td>
<td>0.06(0.07)</td>
<td>-1951.0</td>
</tr>
<tr>
<td>TWW4/EJ</td>
<td>1</td>
<td>7.05</td>
<td>-</td>
<td>45.00</td>
<td>52.05</td>
<td>0.13(0.12)</td>
<td>-993.4</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>7.07</td>
<td>0.000</td>
<td>44.98</td>
<td>52.05</td>
<td>0.14(0.12)</td>
<td>-993.4</td>
</tr>
<tr>
<td>TWWt/EJ</td>
<td>1</td>
<td>14.71</td>
<td>-</td>
<td>137.67</td>
<td>152.33</td>
<td>0.09(0.061)</td>
<td>-3561.2</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>14.72</td>
<td>0.000</td>
<td>135.65</td>
<td>152.32</td>
<td>0.09(0.061)</td>
<td>-3561.1</td>
</tr>
</tbody>
</table>

$\sigma^2_a$: Direct additive genetic variance; $\sigma^2_s$: Permanent environmental variance; $\sigma^2_p$: Residual variance; $h^2_a$: Heritability; $S^2(SE)$: Service sire effect.

in Table 2. Estimations of direct heritability for TWW1/EJ to TWW4/EJ and TWWt/EJ were in the range of 0.03 to 0.14. Direct heritability estimate for TWW4/EJ (0.14) had the highest value among composite reproductive traits. Estimates of the genetic and phenotypic correlations among the composite reproductive traits are presented in Table 3. In general, genetic correlations for composite reproduction traits were positive and moderate to high; and ranged from 0.362 (between second and fourth lambing) to 0.955 (between first and second lambing). High positive genetic correlations (0.959, 0.948, 0.971, and 0.961) were estimated between WW and TWW1/EJ, TWW2/EJ, TWW3/EJ, and TWW4/EJ, respectively.

Estimated co(variance) components and genetic and phenotypic correlations among production traits (BW and WW) and composite reproduction traits (TWW1/EJ to TWW4/EJ) are shown in Table 4. Estimates of (co)variance components, heritability, genetic parameters, and repeatability for Total Weight of lamb Weaned per Ewe Joined from first to fourth lambing (TWWt/EJ) using repeated records model are presented in Table 5.

**DISCUSSION**

Total weight of lamb weaned is a composite reproductive trait with components of lamb growth and survival until weaning and ewe reproductive performance. Despite the low heritability of TWW/EJ in the present study, high coefficient of variance indicates that selection may have the possibility to improve these traits. Also, heritability for composite reproductive traits were in the range of estimates reported by several authors for other breeds (Rosati et al., 2002; Van Wyk et al., 2003; Ekiz et al., 2005; Mokhtari et al., 2010). Also, Vatankhah et
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Table 3. The genetic and phenotypic correlations among the composite reproductive traits. \(^a\)

<table>
<thead>
<tr>
<th>Trait 1</th>
<th>Trait 2</th>
<th>(\sigma_{a,12}^2)</th>
<th>(\sigma_{p,12}^2)</th>
<th>(\sigma_{e,12}^2)</th>
<th>(r_{a,12})</th>
<th>(r_{p,12})</th>
<th>(r_{e,12})</th>
</tr>
</thead>
<tbody>
<tr>
<td>TWW1/EJ</td>
<td>TWW1/EJ</td>
<td>1.549</td>
<td>3.784</td>
<td>5.333</td>
<td>0.955</td>
<td>0.143</td>
<td>0.187</td>
</tr>
<tr>
<td>TWW2/EJ</td>
<td>TWW3/EJ</td>
<td>2.120</td>
<td>3.410</td>
<td>5.530</td>
<td>0.673</td>
<td>0.114</td>
<td>0.166</td>
</tr>
<tr>
<td>TWW3/EJ</td>
<td>TWW4/EJ</td>
<td>1.838</td>
<td>3.377</td>
<td>5.215</td>
<td>0.454</td>
<td>0.114</td>
<td>0.155</td>
</tr>
<tr>
<td>TWW4/EJ</td>
<td>TWW3/EJ</td>
<td>0.344</td>
<td>5.434</td>
<td>6.778</td>
<td>0.674</td>
<td>0.131</td>
<td>0.156</td>
</tr>
<tr>
<td>TWW3/EJ</td>
<td>TWW4/EJ</td>
<td>1.069</td>
<td>7.939</td>
<td>9.008</td>
<td>0.362</td>
<td>0.198</td>
<td>0.204</td>
</tr>
<tr>
<td>TWW4/EJ</td>
<td>TWW4/EJ</td>
<td>2.845</td>
<td>4.735</td>
<td>7.580</td>
<td>0.563</td>
<td>0.102</td>
<td>0.147</td>
</tr>
</tbody>
</table>

\(^a\) \(\sigma_{a,12}^2\): Direct additive genetic covariance of traits 1 and 2; \(\sigma_{r,12}^2\): Residual covariance of traits 1 and 2; \(\sigma_{e,12}^2\): phenotypic covariance of traits 1 and 2; \(r_{a,12}\): Direct additive genetic correlation of traits 1 and 2; \(r_{p,12}\): Residual correlation of traits 1 and 2; \(r_{e,12}\): Phenotypic correlation of traits 1 and 2.

Table 4. The genetic and phenotypic correlations between birth weight or weaning weight and the composite reproductive traits. \(^a\)

<table>
<thead>
<tr>
<th>Trait 1</th>
<th>Trait 2</th>
<th>(\sigma_{a,12}^2)</th>
<th>(\sigma_{p,12}^2)</th>
<th>(\sigma_{e,12}^2)</th>
<th>(r_{a,12})</th>
<th>(r_{p,12})</th>
<th>(r_{e,12})</th>
</tr>
</thead>
<tbody>
<tr>
<td>BW</td>
<td>TWW1/EJ</td>
<td>0.119</td>
<td>0.407</td>
<td>0.526</td>
<td>0.261</td>
<td>-0.201</td>
<td>0.189</td>
</tr>
<tr>
<td></td>
<td>TWW2/EJ</td>
<td>0.423</td>
<td>0.048</td>
<td>0.462</td>
<td>0.572</td>
<td>-0.021</td>
<td>0.124</td>
</tr>
<tr>
<td></td>
<td>TWW3/EJ</td>
<td>0.061</td>
<td>0.329</td>
<td>0.391</td>
<td>0.159</td>
<td>-0.106</td>
<td>0.097</td>
</tr>
<tr>
<td></td>
<td>TWW4/EJ</td>
<td>0.644</td>
<td>0.159</td>
<td>0.484</td>
<td>0.320</td>
<td>-0.775</td>
<td>0.101</td>
</tr>
<tr>
<td>WW</td>
<td>TWW1/EJ</td>
<td>5.095</td>
<td>1.187</td>
<td>3.908</td>
<td>0.959</td>
<td>-0.104</td>
<td>0.219</td>
</tr>
<tr>
<td></td>
<td>TWW2/EJ</td>
<td>0.886</td>
<td>14.797</td>
<td>13.913</td>
<td>0.948</td>
<td>-0.596</td>
<td>0.491</td>
</tr>
<tr>
<td></td>
<td>TWW3/EJ</td>
<td>3.452</td>
<td>1.635</td>
<td>5.082</td>
<td>0.971</td>
<td>-0.083</td>
<td>0.212</td>
</tr>
<tr>
<td></td>
<td>TWW4/EJ</td>
<td>3.225</td>
<td>0.185</td>
<td>3.045</td>
<td>0.961</td>
<td>-0.009</td>
<td>0.125</td>
</tr>
</tbody>
</table>

\(^a\) \(\sigma_{a,12}^2\): Direct additive genetic covariance of traits 1 and 2; \(\sigma_{r,12}^2\): Residual covariance of traits 1 and 2; \(\sigma_{e,12}^2\): phenotypic covariance of traits 1 and 2; \(r_{a,12}\): Direct additive genetic correlation of traits 1 and 2; \(r_{p,12}\): Residual correlation of traits 1 and 2; \(r_{e,12}\): Phenotypic correlation of traits 1 and 2.

Table 5. Estimates of variance components, heritability, and repeatability for TWW/EJ from the first to fourth lambing using repeated records model. \(^a\)

<table>
<thead>
<tr>
<th>Trait</th>
<th>(\hat{G}_{a}^2)</th>
<th>(\hat{G}_{p}^2)</th>
<th>(\hat{G}_{e}^2)</th>
<th>(h^2) (SE)</th>
<th>Pe (SE)</th>
<th>(r)</th>
</tr>
</thead>
<tbody>
<tr>
<td>TWW/EJ</td>
<td>2.234</td>
<td>2.786</td>
<td>27.68</td>
<td>32.701</td>
<td>0.068(0.012)</td>
<td>0.085(0.032)</td>
</tr>
</tbody>
</table>

\(^a\) \(\hat{G}_{a}^2\): direct additive genetic variance; \(\hat{G}_{p}^2\): Permanent environmental variance; \(\hat{G}_{e}^2\): Residual variance; \(h^2\): Phenotypic variance; \(Pe\) : Maternal permanent environmental effect; \(SE\): Standard error; \(r\): Repeatability.

al. (2008) reported that estimation of heritability of total litter weight at weaning (TLWW) was 0.10, which was in the reported range.

In total weight of lamb weaned, phenotypic correlations were less than genetic correlation and varied from 0.147 (between the third and fourth lambing) to 0.513 (between the first and the sum of the four consecutive lambing). High correlations between TWW/EJ in all lambing suggested that selection for TWW in the first lambing would be more useful than later lambing. Also, this indicated that the ability of ewes to produce lambs with high weight at weaning was correlated with WW of the lambs, because the growth of lambs from birth to weaning and mothering ability are affected by individual genotype of lambs (Rosati et al., 2002).

Composite reproductive traits, as combinations of other estimable traits, are expected to have high genetic correlation with each component trait. Genetic correlation of BW with TWW/EJ ranged from 0.159 (with TWW3/EJ) to 0.574 (with
TWW2/EJ) in different lambing and phenotypic correlations among these traits were low and positive (ranging from 0.097 to 0.189). Bromley et al. (2001) reported that genetic correlation between body weight (at birth, weaning and 9th month) and TWW ranged from -0.22 to 0.28, but their definition of total weight of lamb weaned was slightly different from what we used in this study. The low but positive genetic correlations between TWW/EJ and body weights suggest that selection to improve genetic merit in either of these two traits would result in some response in the other trait. High genetic correlations between WW and TWW/EJ in different lambing were reported in previous studies (Snyman et al., 1998b; Olivier et al., 2001). Phenotypic correlations between WW and TWW/EJ in different lambing were lower than genetic correlations and varied from 0.125 (with TWW3/EJ) to 0.491 (with TWW2/EJ). Similar results were obtained by Snyman et al. (1997). Estimates of genetic and phenotypic correlations varied noticeably, but they were within the reported range (Fogarty et al., 1995; Rosati et al., 2002). Positive and high genetic correlations estimated between WW and TWW/EJ were expected, because ewes with more TWW of lambs in each litter would have higher mean WW per lamb at weaning. Several studies have shown that ewe age can affect reproductive traits significantly (Hanford et al., 2006; Vatankhah and Talebi, 2008; Rashidi et al., 2011; Amou Posht-e-Masari et al., 2013). The genetic correlation of composite reproductive traits with BW and WW indicated that selecting ewes based on these traits could improve the other related traits. Estimated phenotypic correlations between production and composite traits were considerably lower than the corresponding genetic correlations.

The estimated fraction of variance due to permanent environmental effects for the Number of Litter Birth (NLB) and Number of Litter Weaning (NLW) were higher than heritability. These estimates indicate that NLB and NLW traits were affected by non-additive genetic effects (dominance and epistasis) and permanent environmental effects (nutrition, management, etc.). Therefore, the first way to improve these traits is to improve environmental effects. The repeatability estimation for composite reproductive traits obtained in our study was 0.153, which was in the range of previously reported studies (Fogarty et al., 1995; Ekiz et al., 2005; Vatankhah et al., 2008). The repeatability estimate was higher than heritability estimates. The accuracy of selection for composite reproductive traits on the first lambing was medium; hence, repeatability can be a measure of correlation between composite reproductive traits in different lambing.

In composite reproductive traits, random effect variance of the service sire was low. In addition, in these traits, the effect of service sire was not significant, so, eliminating this effect had no influence on direct heritability. Increasing trend of the heritability in the third and fourth lambing of TWW/EJ could be due to reduction in increasing environmental variance. That is because during these periods, ewes reached their total body growth and a steady level of milk production; hence, the effect of environment on individuals could be less. The estimated heritability in the third lambing (0.075) was less than that (0.15) in the study of Cloete et al. (2002). Estimated heritability in the fourth lambing (0.135) was similar to that (0.13) of Fogarty et al. (1994).

It has been reported that indirect selection based on early-recorded body weight traits related to TWW could be more effective than direct selection on TWW in Merino composite Afrino sheep (Snyman et al., 1997). On the other hand, indirect selection to increase TWW/EJ based on body weight could increase the genetic merit for reproduction. Although the aim of selection for reproduction is initially to improve the reproductive performance for the present flock, increasing the genetic merit for reproduction of the future generations would be more important.
As the total weight of lambs weaned is affected by several factors such as fertility, litter size, and viability of lambs (Snyman et al., 1997), if the breeding goal is to increase reproductive efficiency, TWW/EJ may be introduced as a selection criterion. On the other hand, moderate to high genetic correlation of TWW/EJ between the first and the other lambing indicates that selection based on this trait in the first period leads to the improvement in other periods, and, consequently, the flock performance may increase. For sheep meat industry in harsh systems, having a large number of heavy lambs at weaning may be the best way to increase performances; hence, selection index including fertility, survival traits, and weight traits can be applied to reach this goal.

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REFERENCES


بهبود صفات تولید ملی و رشد با انتخاب میشها بر اساس صفات تولید مثلی تركیبی در گوسفنده بلوچی

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چکیده

در این مطالعه، اطلاعات 1328 حیوان ایستگاه اصلاح نازگوسفنده بلوچی عباس آباد ایران، بین سالهای 1351 تا 1381 برای برآورد همبستگی های زنیکی و فنوتیپی مجموع وزن بره های شیر گیری TWW2/EJ، TWW1/EJ، TWW3/EJ و TWW4/EJ، و مجموع وزن بره های کل چهار رشذ در زایش های اول، دوم، سوم و چهارم (TWW) و همچنین

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برآورد همیستگی‌های زایش‌های اولتا چهارم با وزن تولد (BW) و شیرگیری (WW) مورد استفاده در مدل‌های DFREML اجراء گردید. این کواریانس همیستگی‌های زایشی تا چهارم با استفاده از نرم افزار TWW/EJ و تا چهارم با استفاده از نرم‌افزار DFREML استفاده گردید. همین‌طور تازهگرفته TWW/EJ و TWW1/EJ در دوره‌های مختلف زایش با وزن تولد و وزن شیرگیری به ترتیب در دامنه‌ی ۱۶/۰ تا ۳۲/۰ و ۹۵/۰ تا ۱۹۷/۰ بودند. نکاتی برای بررسی TWW/EJ در زایش‌های اولتا چهارم ۱۵/۰ بود. نتایج پیش‌نهاد می- که انتخاب برای افزایش عملکرد تولیدی در طول عمر مفید می‌تواند بر اساس در تا چهارم با استفاده از TWW1/EJ به عنوان یک صفت اولین دوره‌ی زایش صورت گیرد. به طور کلی این امر نشان می‌دهد که محدود به جنس شاید بتواند به عنوان معیار انتخاب به صورت غیر مستقیم برای بهبود صفات همیستگی مورد استفاده قرار گیرد.