Effects of Declining Energy Subsidies on Value Added in Agricultural Sector

M. Azamzadeh Shouraki¹*, S. Khalilian¹, and S. A. Mortazavi¹

ABSTRACT

Production subsidies, as a part of the strategy of economic growth of the agricultural sector, are of great importance around the world. Subsidizing production inputs, particularly energy input, is another way of directing subsidy to the agricultural sector. In this research, production function of the agricultural sector was estimated using econometric methods and time series data. After calculating the elasticity of agricultural sector inputs and, simultaneously, estimating their cost and demand functions of production inputs using ISUR (Iterated Simingly Unr elated Regression), farmers’ elasticity of price fluctuation of these inputs was determined. The findings of the production function demonstrated that all inputs, including capital, labor, and energy were used in the optimal production region. The findings of the cost function demonstrated that there was negative and low own elasticity price for inputs, in accord with economic theory. In addition, cross price elasticity of all inputs was positive, i.e. they were substitutes for each other. The findings of the subsidization policy showed that since price elasticity of demand for energy inputs was inelastic, reducing the energy subsidy would reduce energy consumption slightly and, eventually, would decrease value added in the agricultural sector. Finally, it is suggested that the government implements the energy subsidy reduction policy based on cost-benefit analysis.

Keywords: Value added, Energy, Production function, Subsidy.

INTRODUCTION

After more than two decades of general neglect of agriculture as a source of balanced socio-economic growth, there is increasing recognition of the importance of this sector in development. For example, the International Bank for Reconstruction and Development (the World Bank) published its 2008 World Development Report on agriculture entitled “Agriculture for Development”. This publication and its message in support of agriculture as a key sector to economic growth have been welcomed by a wide range of audiences, from politicians to academics (Salami et al., 2012).

Energy markets are important to agriculture. Energy prices affect agricultural production costs directly through fuel and energy use and indirectly through the employment of farm inputs such as fertilizers and chemicals that rely on energy in their manufacturing (Lambert and Gong, 2010). Traditionally, energy is considered as an input for agricultural production. In the new era of renewable energy production, the relationship between energy and agricultural sector has become more interdependent (Aravindhakshan and Koo, 2011).

From the standpoint of the executives of Iran’s economic freedom policy, the low price of agricultural inputs and production subsidies leads to low productivity in the agricultural sector and lower increases in production. The leaders of this policy believe that the low price of production inputs causes suboptimal use of

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these factors. Therefore, they argue that after implementing economic freedom and the consequent increase in the price of inputs, producers would optimize use of production factors (Zibayi and Najafi, 2004).

Debra (2002), in "Agricultural Subsidy in Africa", states the main reasons for supportive policies in the world’s agricultural sector. Those reasons are economic development and growth, especially in rural areas, as well as supporting investment and employment; maintaining domestic product with less dependency on overseas agriculture; and removing or reducing poverty to reach suitable life conditions.

Energy carrying materials have been important ever since they could be used as a substitute for other factors, and have had a significant role in providing goods and services. Energy carrying materials were much used in the manufacturing sector, but technology improvements and machinery for providing services in different sectors, including agriculture, could become important factors in providing goods and services (Fallahi and Khalilian, 2009).

Since agricultural products are essential for provision of a society's needs, governments normally pay subsidies for their production in order to both strengthen the agricultural sector and keep prices low for consumers. Since energy is a strategic input in different sectors of production, its effects on the agricultural sector are studied here.

Subsidy policy is the main supportive tool of government. At the same time, paying agricultural subsidies is one of the main supportive governmental policies benefitting the agricultural sector. By paying subsidies on agriculture inputs, government tries to keep the input price low and to reduce production costs, increase farmers’ incomes, increase domestic producers’ competitiveness and support them, reduce exchange run out, and increase production. On the other hand, farmers may overuse these inputs because of their low price (Shemshadi, 2007).

Commonly, the two inputs of capital stock and labor are used in estimating the total production function. Results of new researches in Iran’s agricultural sector have shown that energy input is very important, next to capital stock and labor inputs, and directly affects the production level (Hozhabr Kiani and Varedi, 2000; Fallahi and Khalilian, 2009). Ma et al. in an article entitled "China's energy economy: Technical change, factor demand and interfactor/interfuel substitution," calculated the Allen partial elasticities of factor and energy substitution, and price elasticities of energy demand for China using a two-stage Translog cost function approach during the study period (1995–2004). The results of their article showed that energy is substitutable for both capital and labor. Coal is significantly substitutable with electricity and complementary with diesel, while gasoline and electricity are substitutable with diesel (Ma et al., 2008).

In a research in Gilan Province, different production functions were estimated and the best one (Translog) was selected, production inputs were calculated, and the Translog cost function was estimated using ISUR and relative and cross elasticity. Allen’s cross elasticity and demand functions of inputs were calculated to investigate the effects of removing the subsidy on fertilizer and insecticides in this province. The findings showed that fertilizer was consumed in Production region II since its demand function was elastic. Removing its subsidy would increase its price and decrease the use of fertilizer in production; moreover, insecticide was consumed in Production region III and was demand elastic. Hence, increasing its price leads to this input being used in Production region II (Azizi, 2005).

Becker (2010) stated that, from 1992 to 2007, the US agricultural market underwent many changes with respect to the inputs they utilized, as well as the prices they paid for them. Among these changes, fuel prices displayed the most severe volatility. By using the Translog Production Function, the price elasticity of substitution was estimated for all agricultural inputs during the period studied, in order to determine how farms change production allocations due to increasing energy prices. It was found that price
elasticities were very low between energy and other inputs, suggesting that farms did not change their input allocations due to increases in energy prices (Becker, 2010).

The present paper addresses the sensitivity of the coefficient of energy input consumption and the value added of the agricultural sector to a one percent change in the energy input price in the agricultural sector. In other words, the present paper explains how sensitive agriculture is to changes in the energy price, based on the policy of energy price increases in Iran.

**MATERIALS AND METHODS**

Determining farmers' sensitivity to price changes is an essential part of policymaking about production subsidies. In this research, a production function was used to determine the agricultural sector's sensitivity to changes in inputs consumption and the sector's sensitivity to inputs' changing prices was determined through cost function and cost share equations.

We investigated the rationality of the agricultural sector by using inputs through the production function, input-output relationship, and the relative importance of every input in increasing the sector's production and determining production elasticity. The consumption amount, which is affected by price changes, and production quantity, which is affected by consumption changes, also are determined through the demand for production inputs.

In this research, in order to estimate the production function and to investigate the long-term and short-term relationships between dependent variables and other dependent variables of pattern, we used an auto regressive distributed lag model (ARDL). Because of the limitations in the methods of the Engel-Granger, Johansen, and Juselius test and error correction model, some research, such as that of Pesaran and Pesaran (1977), tried to find a better method for analyzing long-term and short-term relationships between variables, to cover the shortcomings of these methods. The Pesaran and Pesaran method simultaneously estimates the relationships between the dependent variable and the variables of other dependent models, and solves the problem of variable omission and autocorrelation. Since this model does not have problems such as serial autocorrelation, its estimates are unbiased and effective (Siddiki, 2000).

**Cost Function and Demand Functions**

For the first time, Christensen et al. (1973) studied theoretical principles using the Translog cost function, which is a function of production level and input prices, and is used to estimate the cost function of the agricultural sector and to define inputs' demand functions. The Translog function can be changed using Taylor's Expansion (Carr, 1992):

\[
\ln C = \ln \alpha_0 + \sum_{i=1}^{n} \alpha_i \ln p_i + \frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{n} \gamma_{ij} \ln p_i \ln p_j \\
+ \sum_{i=1}^{n} \gamma_{ii} \ln p_i \ln p_i + \gamma_i \ln y + \frac{1}{2} \gamma_{ii} [\ln y]^2
\]  

(1)

Using the partial derivative of the Translog logarithmic cost function for the \(n^{th}\) input and considering Shepherd's theorem, the demand function of the \(n^{th}\) input becomes:

\[
S_i = \frac{X_i P_i}{C} = \alpha_i + \sum_{j=1}^{n} \gamma_{ij} \ln p_j + \gamma_i \ln y
\]

(2)

Where, \(C = \sum_{j=1}^{n} X_j P_j\) and \(S_i\) is the cost portion of the \(n^{th}\) input.

ISUR was used to estimate the Translog simultaneous equations system and the demand functions of inputs. SUR effectively estimates a group of linear regression equations that are interconnected. Zellner (1962) describes this method. In his work, joint generalized least squares are used to estimate the coefficients of a group of linear regression equations. In estimating equations using ISUR, one of the cost share equations is removed from the system of equation and the other parameters are estimated; then, the parameters of the equation are taken away and the equations are solved based on other parameters. Next, one of the variables is
removed from the cost portion equation, and the price of other inputs is replaced with relative price (relative to the removed input prices) and the total cost is replaced with relative total cost (relative to removed input prices) in the model (Greene, 2003). Since the total cost share equals 1, any item can be removed (Lindert, 1989). However, in the SUR method, the equation that has had the least determination coefficient \( R^2 \) in OLS estimation method is removed. In order to avoid these limitations, ISUR is used instead of SUR, because ISUR estimations are not sensitive to the removed equation of the maximum likelihood (ML) method, which is unique and independent from the removed equations (Barten, 1969). In fact, ISUR repeats the SUR method until reaching convergence.

Accordingly, applying the conditions and assumptions of symmetry and homogeneity in cost function equations, the form of the function changes to:

\[
\text{Relative total cost} = \gamma_j Lny + \frac{1}{2} \gamma_{ij} (Lny)^2
\]  

(3)

\[
S_i = \alpha_i + \sum_{j=1}^{n} \gamma_{ij} Lny + \sum_{j=1}^{n} \gamma_{ij} \frac{P_j}{P_0}
\]  

(4)

\[
A_{ij} = \frac{(S_i + S_j^2 - S_j)}{S_j} \quad i\neq K, L, S
\]  

(5)

\[
A_{ij} = \gamma_{ij} + (S_i, S_j)
\]  

(6)

Based on the relationship between price elasticity of demand and Allen’s substitution elasticity:

\[
E_{ij} = S_j A_{ij}, \quad E_{ii} = S_i A_{ii}
\]  

(7)

The macro data were gathered from the published database of the Central Bank, Energy Balance Sheet, and the Iran Statistics Center (1974–2008).

Consequently, the agricultural production function was estimated and production elasticity of the inputs was calculated. Then the demand function of the agricultural sector’s inputs was estimated to investigate the effect, as well as the effectiveness, of this policy on the value added of the agricultural sector. Next, the way this governmental policy can influence value added was estimated by calculating the price elasticity of demand.

**RESULTS AND DISCUSSION**

In a number of countries including Iran, production functions like the Cobb-Douglas, Transcendental, and Translog are chosen for estimation of the agricultural production function (Torkamani, 1998). In this research, results from investigation of the mentioned functions showed that the Cobb-Douglas function was best suitable for the agricultural sector production function, in agreement with most of the literature on the production function of Iran’s agricultural sector (Amir Teymori and Khalilian, 2008; Soltani, 2005). Cobb-Douglas function’s form is shown below:

\[
\text{LnVA} = C + \beta_1 \text{LnL} + \beta_2 \text{LnK} + \beta_3 \text{LnE} + \sum_{j=1}^{n} \gamma_{ij} \text{LnP}_j + \sum_{j=1}^{n} \gamma_{ij} \text{LnS}_j + D7476
\]  

(8)

The variables of the research are:

\[
\text{LVA}: \text{Natural logarithm of value added in the agricultural sector}
\]

\[
\text{LL}: \text{Natural logarithm of labor in the agricultural sector}
\]

\[
\text{LK}: \text{Natural logarithm of capital in the agricultural sector}
\]

\[
\text{LE}: \text{Natural logarithm of energy in the agricultural sector}
\]

\[
Dj: \text{Imaginary variable related to wartime}
\]

\[
D7476: \text{Imaginary variable related to trend changes in labor of the agricultural sector in 1976.}
\]

The Cobb-Douglas production function was calculated for Iran's agricultural sector...
using Microfit 4.1 and the time series statistics of 1974–2008. To have a better estimation of the Cobb-Douglas function of the agricultural sector in this research, two dummy variables were used to alleviate the structural failure of the agricultural sector capital stock in wartime (Dj) and changes in labor trends before revolution (D7476). The results of the Cobb-Douglas proportion function in the agricultural sector are shown in Table 1.

Since the absolute value of sample t is greater than the critical quantity, \( H_0 \) (null hypothesis, no collective relationship) is rejected, and the long-term equating relationship between the variables in the model is verified. The long-term model coefficient in the Cobb-Douglas function indicates the production elasticity in the function after a 1% change in the dependent variables, i.e., what percentage change will occur in the dependent variable. Production elasticity of labor, capital, and energy is 0.90, 0.29, and 0.60, respectively. (Table 2)

The results of the Cobb-Douglas production function show that the inputs’ production elasticity is between 0 and 1, demonstrating that the inputs are in the optimal area. Overall, we have an increasing return to scale in agricultural sector; therefore, the degree of homogeneity in agricultural sector is greater than one. Therefore, production in this sector can be estimated because, as production increases, the cost of producing each additional unit falls.

As can be seen from Table 2, most of the variables are significant beyond the 5% level. The coefficient of ECM (-1) in the short-term model is -60%. This is statistically significant and negative. A -60% coefficient of ECM means that approximately 60% of value-added variable deviations in the agricultural sector will disappear after one year. In other words, complete justification of the results of implementing a policy requires less than two years’ time.

### Table 1. Testing the existence of a long-run relationship between the variables of the Cobb-Douglas production function pattern.

<table>
<thead>
<tr>
<th>Sample t</th>
<th>Banerjee critical quantity: Dolado and Mestre at 1% level</th>
<th>Banerjee critical quantity: Dolado and Mestre at 5% level</th>
</tr>
</thead>
<tbody>
<tr>
<td>-3.82</td>
<td>-4.59</td>
<td>-6.1</td>
</tr>
</tbody>
</table>

### Table 2. Coefficients of the Cobb-Douglas model.

<table>
<thead>
<tr>
<th>Sig level</th>
<th>Sample t</th>
<th>SD</th>
<th>Coefficient</th>
<th>Variable</th>
</tr>
</thead>
<tbody>
<tr>
<td>[0.080]</td>
<td>1.85</td>
<td>0.49</td>
<td>0.90</td>
<td>LL&lt;sup&gt;a&lt;/sup&gt;</td>
</tr>
<tr>
<td>[0.000]</td>
<td>5.32</td>
<td>0.05</td>
<td>0.29</td>
<td>LK&lt;sup&gt;b&lt;/sup&gt;</td>
</tr>
<tr>
<td>[0.000]</td>
<td>13.34</td>
<td>0.04</td>
<td>0.60</td>
<td>LE&lt;sup&gt;c&lt;/sup&gt;</td>
</tr>
<tr>
<td>[0.000]</td>
<td>14.5</td>
<td>0.2</td>
<td>-4.32</td>
<td>C</td>
</tr>
<tr>
<td>[0.000]</td>
<td>-4.36</td>
<td>0.017</td>
<td>0.39</td>
<td>D7476&lt;sup&gt;d&lt;/sup&gt;</td>
</tr>
<tr>
<td>[0.001]</td>
<td>3.83</td>
<td>0.1</td>
<td>-0.07</td>
<td>Dj&lt;sup&gt;e&lt;/sup&gt;</td>
</tr>
<tr>
<td>[0.000]</td>
<td>-5.63</td>
<td>0.1</td>
<td>-0.60</td>
<td>ECM(-1)&lt;sup&gt;f&lt;/sup&gt;</td>
</tr>
</tbody>
</table>

<sup>a</sup> Natural logarithm of labor in the agricultural sector, 
<sup>b</sup> Natural logarithm of capital in the agricultural sector, 
<sup>c</sup> Natural logarithm of energy in the agricultural sector, 
<sup>d</sup> Imaginary variable related to trend changes in labor of the agricultural sector in 1976, 
<sup>e</sup> Imaginary variable related to wartime, 
<sup>f</sup> The error correction variable in the short-term model
prices, symmetry was estimated for the period of 1974–2008. The results are shown in Table 3.

As can be seen in Table 3, most of the own and cross effects are significant beyond the 5% level. Based on the estimations of cost function, the determination coefficient of the estimated cost function is 0.99, which demonstrates that the independent variables of the cost function strongly describe the dependent variable. This allows us to capture the two elements that tell us a great deal about farming production. These elements are the own-price elasticity and the cross-price elasticities for all inputs. These elasticities will help us to answer how firms in the agricultural sector react to changes in prices. In this research, after estimating the Translog cost function, by using ISUR, own and cross price elasticities of demand were calculated.

Calculating Relative and Cross Price Elasticity of Inputs’ Demand

The cost function of the agricultural sector was used to calculate the relative and cross price elasticity of inputs’ demand. The results are shown in Table 4.

The findings show that each of the own price elasticities of inputs is negative, in accord with economic theories. This implies that increasing the price of each input decreases the demand for it. The absolute value of price elasticity of demand for all inputs (capital, labor, and energy) is less than 1, meaning that these inputs are inelastic to changing price. In other words, 1% change in input price leads to less than a 1% change in input demand. Cross price elasticity of demand between energy to capital and energy to labor is 0.088 and 0.033, respectively. This means that increasing the energy price increases the

<table>
<thead>
<tr>
<th>Parameter</th>
<th>Estimated value</th>
<th>Sample t</th>
<th>Parameter</th>
<th>Estimated value</th>
<th>Sample t</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\alpha_0$</td>
<td>1178.139*</td>
<td>6.76</td>
<td>$\gamma_{ee}$</td>
<td>-0.05</td>
<td>–</td>
</tr>
<tr>
<td>$\alpha_k$</td>
<td>2.93*</td>
<td>11.78</td>
<td>$\gamma_{lk}$</td>
<td>-0.009*</td>
<td>-2.29</td>
</tr>
<tr>
<td>$\alpha_l$</td>
<td>-0.54**</td>
<td>-2.3</td>
<td>$\gamma_{le}$</td>
<td>0.0005</td>
<td>–</td>
</tr>
<tr>
<td>$\alpha_e$</td>
<td>-1.39</td>
<td>–</td>
<td>$\gamma_{ke}$</td>
<td>-0.032</td>
<td>–</td>
</tr>
<tr>
<td>$\gamma_y$</td>
<td>-75.56*</td>
<td>-12.5</td>
<td>$\gamma_{ly}$</td>
<td>0.011</td>
<td>1.57</td>
</tr>
<tr>
<td>$\gamma_{yy}$</td>
<td>1.23*</td>
<td>12.73</td>
<td>$\gamma_{ky}$</td>
<td>-0.072*</td>
<td>-2.96</td>
</tr>
<tr>
<td>$\gamma_{kk}$</td>
<td>0.041*</td>
<td>11.30</td>
<td>$\gamma_{ey}$</td>
<td>0.061</td>
<td>–</td>
</tr>
<tr>
<td>$\gamma_{ll}$</td>
<td>0.009*</td>
<td>5.85</td>
<td>$\gamma_{ly}$</td>
<td>0.012</td>
<td>0.015</td>
</tr>
<tr>
<td>$\gamma_{kl}$</td>
<td>0.25</td>
<td>-0.012</td>
<td>$\gamma_{ly}$</td>
<td>0.087</td>
<td>–</td>
</tr>
<tr>
<td>$\gamma_{le}$</td>
<td>0.088</td>
<td>0.034</td>
<td>$\gamma_{ly}$</td>
<td>-0.096</td>
<td></td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Variable</th>
<th>Capital</th>
<th>Labor</th>
<th>Energy</th>
</tr>
</thead>
<tbody>
<tr>
<td>Capital</td>
<td>-0.02</td>
<td>0.012</td>
<td>0.015</td>
</tr>
<tr>
<td>Labor</td>
<td>0.25</td>
<td>-0.012</td>
<td>0.087</td>
</tr>
<tr>
<td>Energy</td>
<td>0.088</td>
<td>0.034</td>
<td>-0.096</td>
</tr>
</tbody>
</table>
Table 5. The effects of subsidy policy on value added in the agricultural sector.

<table>
<thead>
<tr>
<th>Input</th>
<th>Capital</th>
<th>Labor</th>
<th>Energy</th>
</tr>
</thead>
<tbody>
<tr>
<td>Price elasticity of Demand</td>
<td>-0.02</td>
<td>-0.012</td>
<td>-0.096</td>
</tr>
<tr>
<td>Production elasticity</td>
<td>0.29</td>
<td>0.9</td>
<td>0.6</td>
</tr>
<tr>
<td>The percentage change in value added</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>of the agricultural sector because</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>of a 1% change in input price</td>
<td>-0.0058</td>
<td>-0.01</td>
<td>-0.057</td>
</tr>
</tbody>
</table>
Moreover, since labor production elasticity is 0.90, it can be inferred that decreasing the labor in agricultural sector by 1% leads to a 0.90% reduction in value added. Finally, by combining price elasticity of demand and production elasticity of labor in the agricultural sector, a 1% increase in the labor process increases value added in the agricultural sector by 0.01%. Consequences of subsidy policy on value added in the agricultural sector are shown in Table 5.

Analyzing Profits and Losses of a Subsidy Reduction Policy

In reducing the energy subsidy of the agricultural sector, government should consider the consequent profits and losses. Government should implement this policy only if its benefits cover its losses. The profits and losses of implementing energy subsidy reduction are shown in Table 6.

The indirect profits of reducing the energy subsidy are pollution reduction, reducing government's financial burden, and the responsibility for maintaining and distributing energy.

CONCLUSIONS

Based on our research and the preceding calculations, we conclude that:

1. The production elasticity of every

Table 6. Comparison of the benefits and losses of a 1% increase in energy price.

<table>
<thead>
<tr>
<th>Profit</th>
<th>Loss</th>
</tr>
</thead>
<tbody>
<tr>
<td>1. Reducing energy consumption by about 0.05% is worth X</td>
<td>1. Reducing the value added of the agricultural sector because of using less inputs</td>
</tr>
<tr>
<td>2. Saving the energy subsidy</td>
<td></td>
</tr>
</tbody>
</table>
production input is between zero and 1, which means these inputs are economically in the second production region. Since all of the inputs, energy in particular (which is the main subject of this research), are used in the optimal production region, it is suggested that government implement this policy stage by stage, because a sudden increase in prices may bring a price shock to the agricultural sector and cause more damage and deterioration to the agricultural sector.

The price elasticity of demand for all inputs is low, which means that increasing the energy price by 1% decreases the use of that input less than 1%. Consequently, energy demand is inelastic. If government increases the energy price by 1%, the agricultural sector value added decreases by 0.057%.

Low elasticity of demand for production inputs shows that price changes have negligible effect on the use of these inputs. It shows that farms will not change their input allocations due to increases in energy prices.

Based on positive cross elasticity of demand between every two inputs, substitution among inputs is confirmed. Therefore, if the price of each one increases, the other one should replace it.

Positive cross elasticity of demand between energy and labor indicates the possibility of their substitution. If energy price is increased, experienced labor can take the place of machines that are high consumers of energy; and vice versa.

Positive cross elasticity of demand between energy and capital indicates the possibility of their substitution, which means that when the energy price increases, using advanced machines to reduce energy consumption is preferred. Therefore, when the energy price increases, government can give some debt capital to farmers to help them buy modern machines and indirectly prepare for increasing productivity and reducing energy consumption.

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بکری اثر کاهش پارانه انرژی بر ارزش افزوده بخش کشاورزی

م. اعظم زاده، ص. خلیلیان، و س. غ. مرتضوی

چکیده

بکری اثر کاهش پارانه انرژی بر ارزش افزوده بخش کشاورزی، پارانه‌های تولیدی در کشورهای جهان از جایگاه خاصی برخوردار است و یکی از راه‌های برداشت پارانه‌های تولیدی در بخش کشاورزی اعطای پارانه به نهاده‌های تولیدی، از جمله نهاده انرژی می‌باشد. در این تحقیق با استفاده از روشهای اقتصاد منتج، و با استفاده از اطلاعات سری زمانته، ابتدا تابع تولید بخش کشاورزی تخمین زده شد و پس از محاسبه خشکاله‌های تولید نهاده‌های بخش کشاورزی، با تخمین هم‌مان ارزش تولید نهاده‌های انرژی و تفاوت‌های نهاده‌های ISUR تولیدی به وسیله حساب کشاورزان نسبت به نیازهای قائم این نهاده‌ها منشأشده است. نتایج بدست آمده از تابع تولید بخش کشاورزی نشان داد که هر سه نهاده سرمایه، نیروی کار و انرژی در
ناتجی بهبود تولید مصرف می‌شود. نتایج بدست آمده از تابع هزینه نشان داد که کلیه قیمت خودی نهاد‌ها، منفی و کم کلیه است که سازگار با تنوری اقتصاد می‌باشد. همچنین کلیه قیمتی متقابل تمامی نهاد‌ها باید می‌باشد که بیانگر جانشینی هر سه نهاد به جای یکدیگر می‌باشد. نتایج حاصل از سیاست پاره‌ای نشان داد که کلیه یارانه‌های نهاد مصوب به دلیل کلیه ناپذیر بودن نقاشی این نهاد نسبت به قیمت باعث کاهش اندازه در مصرف آن نهاده گردد و در نهایت ارزش افزوده بخش کشاورزی را کاهش می‌دهد. در پایان پیشنهاد می‌شود دولت با توجه به تحلیل هزینه- فایده اقدام به اجرای سیاست کلیه یارانه انرژی نماید.